



Genetic variants underlying differences in facial morphology in East Asian and European populations

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Facial morphology—a conspicuous feature of human appearance—is highly heritable. Previous studies on the genetic basis of facial morphology were performed mainly in European-ancestry cohorts (EUR). Applying a data-driven phenotyping and multivariate genome-wide scanning protocol to a large collection of three-dimensional facial images of individuals with East Asian ancestry (EAS), we identified 244 variants in 166 loci (62 new) associated with typical-range facial variation. A newly proposed polygenic shape analysis indicates that the effects of the variants on facial shape in EAS can be generalized to EUR. Based on this, we further identified 13 variants related to differences between facial shape in EUR and EAS populations. Evolutionary analyses suggest that the difference in nose shape between EUR and EAS populations is caused by a directional selection, due mainly to a local adaptation in Europeans. Our results illustrate the underlying genetic basis for facial differences across populations.

Facial morphology has substantial variations at the individual and population level. Multiple studies show substantial differences in craniofacial morphology across people from different geographic regions^{1–18}. For example, individuals of European ancestries (EUR) have a more protruding nose and brow ridges than those of East Asian ancestries (EAS)¹⁹. Such differences must have a strong genetic basis, which remains unknown due largely to the low number of studies performed in EAS populations compared with EUR populations. Previous genetic studies collectively reported about 219 loci associated with facial morphology in EUR populations, but only 24 were reported in EAS or Eurasian-ancestry populations, about 20 in Latin American-ancestry populations and 4 in African-ancestry populations (Supplementary Table 1)^{3–5,9,12–14,16,17}. Large-scale studies in EAS populations and other non-European populations are much needed to provide a complete architecture of the genetic basis of facial morphology, particularly the observable differences across populations.

Here, we performed a genome-wide association study (GWAS) based on a large collection of three-dimensional (3D) facial images from the Han Chinese population. Using a data-driven phenotyping approach, we identified hundreds of associated variants^{10,18}. In addition, we identified specific variants distinguishing facial appearance

between EUR and EAS populations. We further provided evidence that those population-based facial differences, especially nose shape, were under selection. A schematic overview of our study design can be found in Extended Data Fig. 1.

Results

GWAS on facial phenotypes discovered 244 leading variants. To study facial variation from a global to a local scale, we used 3D facial surface scans from a large-scale EAS population (Supplementary Table 2; Methods) in a discovery ($n=6,968$) and replication ($n=2,706$) cohort, and subsequently combined them in a meta-analysis. A semisupervised phenotyping procedure defined 63 hierarchically arranged facial segments using the discovery cohort (Methods). Next, we performed a canonical correlation analysis (CCA)-based GWAS on each facial segment's group of principal components (Methods). Subsequently, we identified 50 independent tests using parallel analysis and permutation test (Supplementary Note; Methods)^{18,20,21}. Thus, besides conventional genome-wide significant threshold ($P=5\times 10^{-8}$), we set a stricter study-wide significant threshold to $P=9.8\times 10^{-10}$ ($P=5\times 10^{-8}/51.41$) after Bonferroni correction for multiple testing. In the discovery datasets, we identified 153 genome-wide significant

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variants ($P < 5 \times 10^{-8}$, minor allele frequency (MAF) > 0.05) in 124 loci after condition analysis and peak selection (Methods). In the replication dataset, 119 out of 153 (77.8%) variants were replicated at nominal significance ($P < 0.05$), 118 (77.1%) were replicated at a false discovery rate (FDR) < 0.05 , 75 (49.0%) were replicated at a Bonferroni-corrected significance of $P < 3.28 \times 10^{-4}$. For the 84 variants that passed the study-wide significance ($P < 9.8 \times 10^{-10}$), 80 out of 84 (95.2%) were replicated at nominal significance ($P < 0.05$) and FDR < 0.05 , 65 (77.4%) were replicated at a Bonferroni-corrected significance of $P < 5.95 \times 10^{-4}$. To increase statistical power, we performed a meta-analysis using Stouffer's method to combine the P values obtained from the discovery and replication cohort^{18,22}. As a result, we identified 244 independent variants in 166 loci under the genome-wide threshold associated with normal-range facial variation (Fig. 1a and Supplementary Table 3) where 151 variants in 106 loci were study-wide significant. According to the anatomical structure, we classified the 244 genome-wide significant variants into ten facial regions (Supplementary Note), including forehead, glabella, eye, tempora, zygoma, nose, maxillary, upper mouth, lower mouth and mandible (Fig. 1b and Supplementary Table 4). The nose was the feature associated with the most variants (107) out of the ten regions. The numbers of variants associated with the other nine regions were: glabella (35), upper mouth (32), eye (29), zygoma (28), maxillary (25), mandible (20), forehead (16), lower mouth (13) and tempora (12). In addition, we calculated the genome-wide heritability of each 3D facial segment at each level. The heritability of 3D facial segments ranged from 7.47% to 52.3%. As we expected, the nose segments were also among the most heritable area (42.17% to 46.86%; Supplementary Fig. 1).

We considered a variant as new when it was not in linkage disequilibrium ($LD, r^2 < 0.1$) with previously reported variants in facial GWASs ($P < 5 \times 10^{-8}$; Supplementary Table 1) in any EAS, EUR and African (AFR) populations in the 1000 Genomes Project Phase 3 (1000GP; Methods)²³. As such, 130 of the 244 leading variants under the genome-wide threshold were new, while 65 of the 151 under the study-wide threshold were new (Supplementary Table 3). We considered a genetic locus as new when it did not overlap with previously reported genomic loci associated with facial variation. As such, 62 out of 166 loci were new.

We used FUMA and GREAT to annotate the leading variants^{24,25}. As a result, we identified 206 candidate genes potentially associated with facial variation, among which 100 genes were not reported in previous facial GWASs (Supplementary Table 3). We found that the genes associated with the leading variants were highly enriched in biological processes of skeletal system development and morphogenesis (Extended Data Fig. 2a). Moreover, the epigenome and transcriptome datasets showed that the leading variants were enriched mainly for enhancers in craniofacial tissues. Compared with the ectoderm at the later stage of fetal development, candidate genes were highly expressed in the mesenchyme (Extended Data Fig. 2b,c), consistent with our expectation and previous studies^{18,26–30}.

We investigated associations ($P < 1 \times 10^{-5}$) with other complex traits for the 244 variants through PhenoScanner, a web-based GWAS repository (Methods)³¹. We found that the traits sharing high genetic components with facial shapes involved mainly physical measurements, body composition and hair morphology (Supplementary Table 5). The highest coassociation was the risk of atrial fibrillation (*PITX2*, rs6843082; $P = 3.0 \times 10^{-155}$). This variant has been reported to be associated with cardioembolic stroke and ischemic stroke, suggesting that facial features might be a biomarker of cardiovascular disease.

Characteristics of specific variants in EAS and EUR. By comparing the 244 leading variants identified in our study with the 203 leading variants reported in a recently published EUR study using similar phenotyping and analysis framework, 89 variants were

shared in both studies (Methods)¹⁸. The remaining 155 and 114 variants in the EAS and EUR studies respectively were different (Fig. 2a and Supplementary Table 6). Therefore, we defined three different groups: 89 shared variants, 155 EAS-specific variants and 114 EUR-specific variants.

To understand these shared and population-specific variant characteristics, we examined their allele frequencies in EAS and EUR populations based on the 1000GP²³. Comparison of crosspopulation MAF showed that the groups of population-specific variants had higher MAF in their respective population ($P_{\text{EAS-specific}} = 9.0 \times 10^{-12}$, $P_{\text{EUR-specific}} = 8.1 \times 10^{-9}$). In contrast, the shared variants had no difference of MAF between the two populations ($P_{\text{shared}} = 0.1$) (Fig. 2b–e). These results suggest that higher MAFs may increase the statistical power to detect the variants associated with facial variation in the respective populations. Moreover, 77.5% (69 out of 89) of the shared variants passed the study-wide significance threshold while only 54.2% (84 out of 155) EAS-specific variants and 55.2% (63 out of 114) EUR-specific variants passed the study-wide significance threshold (chi-squared test $P_{\text{EAS}} = 3.2 \times 10^{-4}$; $P_{\text{EUR}} = 4.8 \times 10^{-6}$).

To explore potential biological functional differences between the shared and population-specific variants, we used Metascape to compare the differences of enriched terms for their annotated genes³². We found that EAS-specific, EUR-specific and shared genes were all associated with top terms that were previously reported in association with craniofacial variation (Fig. 2f). Moreover, we observed a considerable number of functional overlaps among the three groups (Fig. 2g). These results indicate that the associated genes across populations share substantially similar biological processes. We next compared epigenetic regulation patterns of the shared and population-specific variants in various cell types or tissues. The shared and population-specific variants were all enriched for enhancers in craniofacial tissues (Fig. 2h), again indicating that the potential functions of facial variants across populations are analogous.

As expected, we found that the population-specific variants had a significantly higher Wright's fixation index (F_{ST}) than random variants from a genome-wide background both in EAS and EUR populations ($P_{\text{EAS-specific}} = 2.42 \times 10^{-10}$; $P_{\text{EUR-specific}} = 0.0063$; Extended Data Fig. 3), while the shared variants had no significant difference compared with random variants ($P_{\text{EAS-shared}} = 0.599$)^{33,34}. The same result applied to the crosspopulation extended haplotype homozygosity (XP-EHH) analysis using REHH2 in these two populations ($P_{\text{EAS-specific}} = 0.0078$; $P_{\text{EUR-specific}} = 0.038$; $P_{\text{EAS-shared}} = 0.449$)^{35,36}. These results suggest that facial variation across populations could be attributed to random drift and natural selection. The population-specific variants explained a larger proportion of natural selection while the shared variants may mainly explain random drift influencing facial variation.

In summary, the population-specific variants found in EAS and EUR were identified due to crosspopulation MAF differences and subtle effect sizes. Nonetheless, similar biological processes underlying facial variation were shared across populations.

Polygenic shape analysis generalizes results from EAS to EUR.

To explore the genetic basis of EAS facial shape and the genetic factors contributing to the difference in facial shape between EAS and EUR, we first investigated whether the association results of the leading variants in our EAS study could be generalized to EUR.

We introduced a new polygenic shape analysis to investigate whether the differential accumulated genetic effects between the two populations of the leading variants is in line with the actual population facial differences. Similar to the classic polygenic score analysis, we defined the polygenic shape (PS) for an individual as the sum of the number of effect risk alleles weighted by risk allele effect size of all the leading variants³⁷. Similarly, the Polygenic Population Shape (PPS) is the average polygenic shape for a given population

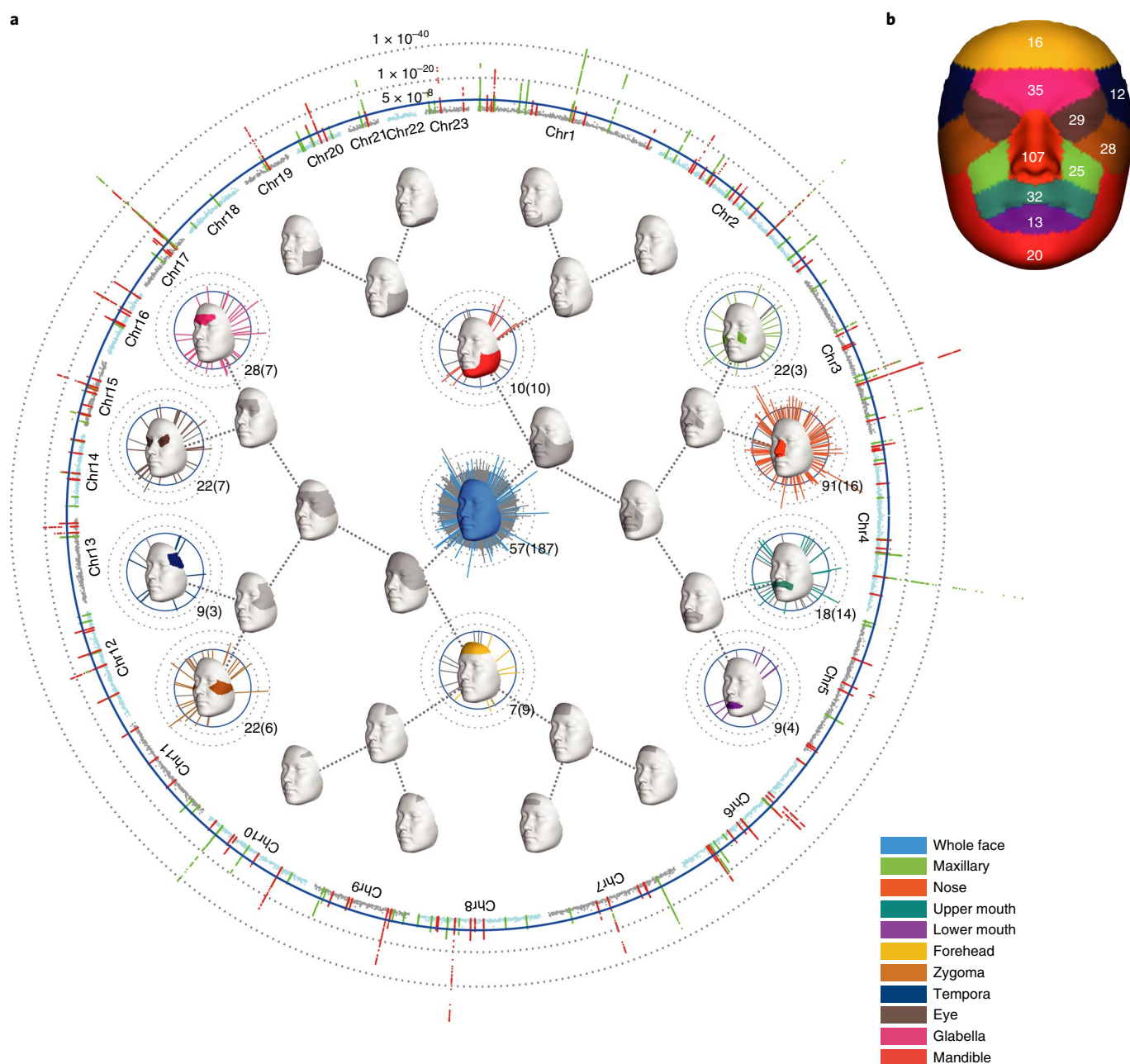


Fig. 1 | Overall results of genome-wide association meta-analyses in Han Chinese cohort. a, Manhattan plot on the ring represents the meta-analyses P values, with chromosomes colored and labeled. P values are $-\log_{10}$ scaled. Using 500 kb windows, peak variants are colored in red (new variants) and green (known variants). Inside the Manhattan plot, a binary tree of facial segments illustrates the hierarchical facial segmentation up to level four. On the whole face and ten selected facial segments (colored), variants in the segment itself (colored) and in its derivative segments (gray) are plotted, with the number of variants listed (number of variants in derivative segments in parenthesis). **b**, Visualization of ten selected facial segments and number of significant variants.

(Methods). Using data of EUR ($n=404$) and EAS ($n=208$) individuals from 1000GP, we calculated the PPS of the two populations for the whole face and ten anatomical facial regions.

To visualize and compare this effect with the true population facial shape average for each facial region, we constructed EUR and EAS PPS-derived faces by adding and subtracting, respectively, $(PPS_{EUR} - PPS_{EAS})/2$ to a population neutral average face, which was constructed as the average of the EUR and EAS population average shapes (Methods). We used 3D facial scans of EAS and EUR individuals to calculate each population's average face and therefore generated EUR and EAS average faces. Compared with the average face of EUR, we found that EAS had more protrusion in the cheek;

more concavity in the forehead, glabella, nose and mandible (Fig. 3a,i), which were consistent with a previous study¹⁹. Interestingly, when we amplified the differentiated accumulated genetic effects five times, the PPS-derived faces looked very similar to EAS and EUR's actual average face (Fig. 3a,ii). The EUR and EAS PPS-derived faces showed similar facial variation to the ground truth, especially in the glabella and nose region (Fig. 3a,ii).

To test the generalization of the association results from our study to EUR, we compared the EUR and EAS PPS-derived faces of the whole face using all 244 leading variants with the PPS of the whole face using 244 variants chosen randomly from the genome. The EAS and EUR PPS-derived faces using the leading variants

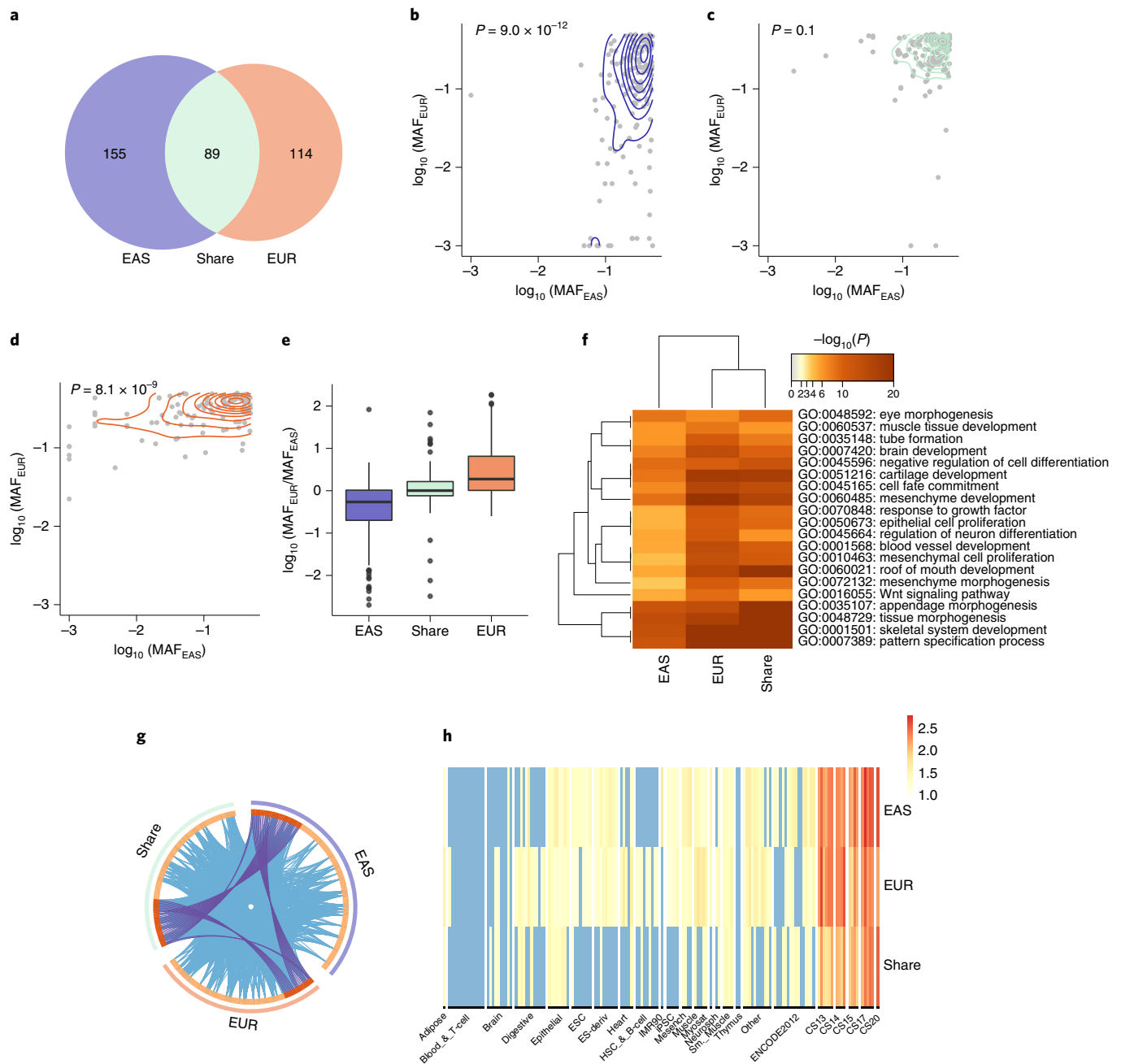


Fig. 2 | Comparison of shared and population-specific variants. a, Number of specific and shared variants of two study cohorts (EAS and EUR). **b–d**, Crosspopulation MAF comparisons of specific variants of EAS ($n=155$) (**b**), shared variants ($n=178$ (89 in EAS study and 89 in EUR study; shared variants in both studies are used) (**c**) and specific variants of EUR ($n=114$) (**d**). In **b–d**, P values are provided using a two-sided Mann–Whitney U -test. When $\text{MAF} < 0.001$, MAF was truncated to 0.001 to fit the log scale (1000GP). **e**, MAF ratio comparison of three groups. Boxplots show medians (center lines), first and third quartiles (lower and upper box limits, respectively), 1.5-fold interquartile ranges (whisker extents) and outliers (black circles). Colors as in **a**. **f**, Metascape analysis shows the biological processes associated with genes in the three groups of genes. **g**, Each outside arc represents a group and each inside arc represents a gene list. On the inside, each arc represents a gene list, where each gene has a spot on the arc. The dark orange color represents overlapping genes among groups. Same genes (purple lines) and different genes fell into the same ontology term (blue lines) of three groups. **h**, Heatmap indicates the global enrichment of trait-associated variants of each group (y axis) in enhancer of different tissue (x axis).

were significantly more similar to the true population average faces than when using random variants, as measured either by Euclidean distance ($P_{\text{dis}} = 0.007$) or cosine similarity ($P_{\text{EUR}} = 0.004$; $P_{\text{EAS}} = 0.005$; Fig. 3b). Moreover, we calculated individual facial polygenic shape for EUR and EAS individuals from 1000GP and measured their East Asian ancestry facial appearance (EAS-FA), defined as the projected length of the individual's polygenic shape onto the explicit

EAS–EUR shape difference (Supplementary Fig. 2; Methods). The EAS-FA of both EUR and EAS groups were significantly separated, with EUR individuals closer to the EUR face and EAS individuals closer to the EAS face (t -test $P < 2.2 \times 10^{-16}$; Fig. 3c).

We also performed the same analyses locally for ten anatomical facial regions. The EAS and EUR PPS-derived shapes using the leading variants were more similar to the population average shapes

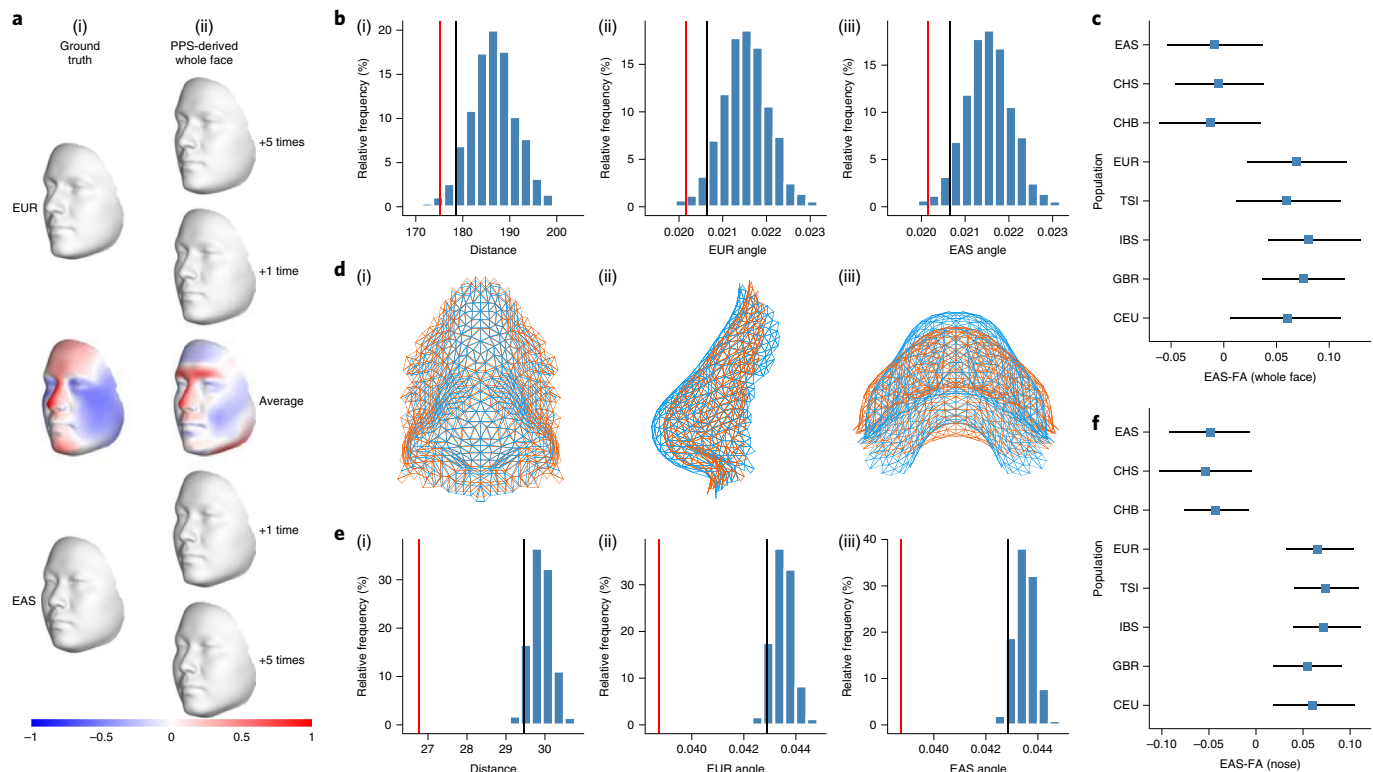


Fig. 3 | Visualization of PPS-derived whole faces and nose of EAS and EUR, and statistical validation of the PPS approach. **a**, Visualization of facial morphology of EAS and EUR. (i) from top to bottom, representing the ground truth of EUR average face, overall average face (EUR and EAS) and EAS average face respectively; (ii) from top to bottom, the PPS-derived whole faces of 244 leading variants by adding $PPS_{EUR} - PPS_{EAS} +5, +1$ to $-1, -5$ times on the overall average facial shape. Differences are visualized using the normal displacement (displacement in the direction locally perpendicular to the facial surface); blue and red refer to depression and protrusion in local shape, respectively. **b**, The null distribution (blue) of (i) Euclidean distance, (ii) cosine similarity with EUR average face and (iii) cosine similarity with EAS average face using 1,000 simulations from 244 random variants on the whole face; the red line infers the statistics of the 244 leading variants and the black line infers 95% quantile of distribution from the random variants. **c**, The EAS-FA of polygenic shapes (whole face) for individuals in 1000GP. The squares represent the mean EAS-FA score and the horizontal lines represent the first and third quantile. **d**, The $PPS_{EUR} - PPS_{EAS}$ difference of the nose region in three views. (i), (ii) and (iii) indicate front, side and vertical views, respectively. The PPS-derived nose of EUR and EAS are presented in the blue and orange color, respectively. **e**, The distribution of (i) Euclidean distance, (ii) cosine similarity with EUR average nose and (iii) cosine similarity with EAS average nose using 1,000 simulations from random variants; the red line infers the statistics of the 107 leading variants associated with nose and the black line infers 95% quantile of distribution from the random variants. **f**, The EAS-FA of polygenic shapes (nose only) for individuals in 1000GP. The squares represent the mean EAS-FA score for the nose and the horizontal lines represent first and third quantile.

than those using random variants, involving the upper mouth, nose, maxillary, glabella, eye, tempora and zygoma in all three measurements of similarity, including cosine similarity, Euclidean distance and EAS-FA (Extended Data Fig. 4a). Notably, the nasal region performed the best among the ten anatomical regions ($P_{dis} = 4.8 \times 10^{-33}$; $P_{EUR} = 3.7 \times 10^{-37}$; $P_{EAS} = 3.4 \times 10^{-37}$; Fig. 3d–f). However, we obtained nonsignificant results for the mandible, forehead and lower mouth (Extended Data Fig. 5). These results indicate that the PPS constructs morphological variations in most facial regions between EUR and EAS. Our results demonstrate that PPS-derived faces using leading variants are similar, both visually and statistically, to the true population average shapes at global and local scales, suggesting that the facial shape effects identified from the EAS study generalize well to EUR populations.

Variants contribute to EAS-FA. Among the leading variants, we aimed to find those that confer more East Asian facial features on EAS, in other words, those variants that increase EAS-FA. We defined the contribution of a variant to EAS-FA as the effect allele frequency weighted projected length of its effect size vector onto the EAS–EUR shape difference (Supplementary Figs. 3 and 7;

Methods). A variant with a positive EAS-FA contribution may cause the facial morphology of the EAS population to increase EAS-FA. In contrast, a variant with a negative EAS-FA contribution may cause the EAS population to increase EUR facial appearance. In brief, variants with large positive EAS-FA are of interest in this context. In each facial region, we constructed a distribution of the contribution of the variant to EAS-FA using 244 leading variants. We further calculated whether each leading variant makes a significantly higher contribution to EAS-FA than the distribution after Bonferroni correction ($P < 0.005$; Extended Data Fig. 6 and Supplementary Table 8). As a result, 13 variants that passed filtering were considered to increase EAS-FA (Table 1 and Supplementary Fig. 4).

The 13 variants had a higher F_{ST} than the other variants (t -test $P < 1.0 \times 10^{-16}$), indicating that they have significant allele frequency differences between EUR and EAS populations. The Population Branch Statistics (PBS) values of the variants were significantly higher in EAS ($P < 1.0 \times 10^{-16}$) relative to EUR and YRI (Yoruba), but not in EUR ($P = 0.188$) relative to EAS and YRI, which suggests that these variants may be under selection in EAS³⁸. Thus, these variants potentially contribute to the morphological differences between EUR and EAS. Most EAS-FA variants might be standing genetic

Table 1 | The 13 variants associated mainly with EAS-FA

rsID	CytoBand	A1	A2	P value	EUR	EAS	PBS_ EAS	PBS_ EUR	F_{ST} _ EURvEAS	Segments	Candidate gene	Allele age
rs7516137	1p36.32	C	G	9.75×10^{-29}	0.318	0.553	0.077	0.036	0.107	18	<i>PRDM16</i>	537,708
rs6669519 ^{ac}	1p31.1	T	A	3.40×10^{-8}	0.173	0.781	0.831	0.000	0.547	24	<i>LHX8</i>	51,373
rs10923710	1p12	G	T	1.20×10^{-44}	0.193	0.507	0.262	0.000	0.207	19, 26	<i>TBX15</i>	88,185
rs3827760	2q12.3	A	G	2.17×10^{-13}	0.000	0.921	2.587	0.306	0.945	27	<i>EDAR</i>	36,410
rs12473319 ^{abc}	2q37.1	G	C	1.35×10^{-10}	0.022	0.478	0.554	0.161	0.511	24	<i>DIS3L2</i>	39,385
rs12632544	3q23	T	A	1.87×10^{-65}	0.000	0.500	0.571	0.332	0.595	24, 25, 26	<i>MRPS22</i>	626,678
rs147468294	6q14.3	A	AC	9.02×10^{-17}	1.000	0.690	0.299	0.207	0.397	7	<i>TBX18</i>	46,640
rs111847181	8p23.1	G	GAC	5.28×10^{-9}	0.454	0.964	0.779	0.000	0.424	18	<i>PPP1R3B</i>	788,343
rs4749259 ^{abc}	10p12.1	T	C	3.88×10^{-29}	0.936	0.584	0.370	0.037	0.334	26	<i>MKX</i>	298,528
rs12258832 ^{abc}	10p12.1	A	G	1.61×10^{-24}	0.892	0.690	0.133	0.005	0.129	26	<i>MKX</i>	64,715
rs3740550 ^a	10q26.11	A	G	6.70×10^{-43}	0.994	0.875	0.127	0.029	0.145	19	<i>RAB11FIP2, EMX2</i>	64,603
rs8068343 ^a	17q24.3	C	T	3.32×10^{-51}	0.959	0.462	0.470	0.281	0.528	18	<i>SOX9</i>	1,261,665
rs9980535 ^{abc}	21q21.3	A	G	3.99×10^{-11}	0.176	0.762	0.357	0.380	0.521	18	<i>LINC00161</i>	1,030,805

^aNew variants in our GWAS finding, which are not in LD ($r^2 < 0.1$) with variants reported in previous facial GWAS studies, see Supplementary Tables 1, 3. ^bNew loci in our GWAS finding, which are not in the same loci (<500 kb) in previous facial GWAS studies, see Supplementary Tables 1, 3. ^cNew genes in our GWAS finding, which are not reported in previous facial GWAS studies.

variations, as the alternative allele frequency was relatively high, given the evolutionary time, as shown in Table 1. Notably, six variants had F_{ST} above 0.5 between EUR and EAS. Furthermore, three of these variants affected the glabella segments and two affected the nasal region. This result suggests that local adaptation might play a role in forming facial variation between EAS and EUR. Among the 13 variants, 6 were reported to be associated with facial shape variation. Well-known facial genes such as *EDAR*, *TBX15* and *MRPS22* have been associated with craniofacial morphology in many studies^{3,4,6,10,13,15,16,18,39–43}. Our study showed a variant in an intron of *TBX15* (rs10923710, NC_000001.10:g.119502774G>T) contributing to maxillary and tempora shape in EAS, consistent with the observation that this locus had multiple spatial effects on the face¹⁸. The rs12632544 (NC_000003.11:g.138946868T>A) is an intergenic variant near *MRPS22* on chromosome 3q23. It is in LD ($r^2 = 0.932$) with rs12633011, which was reported to be associated with eye morphology in a previous EAS study¹⁶. *MRPS22* was reported to be associated with human earlobe size and a mouse skeleton phenotype^{3,44,45}. A reanalysis of a GWAS study on cranioskeletal variation in outbred mice showed that variants in the region of chromosome 9, overlapping with *Mrps22*, were significantly associated with craniofacial variation (FDR <5%; Supplementary Fig. 5). These variants were associated with protrusion of the maxillary bone, and shrinkage of the eye and malar bone. In our study, rs12632544 contributes to EAS-FA in the glabella, eye and tempora (Extended Data Fig. 4b), in line with the result in outbred mice. We also identified seven new variants contributing to EAS-FA, five which are located close to previously unknown genes. Some of these have been reported in the context of craniofacial dysmorphology. For instance, rs6669519 (NC_000001.10:g.75584009T>A), which contributes to the shape of glabella, is an intergenic variant near *LHX8* on chromosome 1p31.1. *LHX8* (LIM Homeobox 8) was reported to be associated with cleft palate, forebrain neuron development and differentiation^{46,47}. The variant rs12473319 (NC_000002.11:g.232880971G>C)—an intronic variant of the *DIS3L2* gene—showed association with EAS-FA in glabella (Extended Data Fig. 4c). In a mouse genome study, *DIS3L2* was found to affect skeleton phenotype^{45,48}. In addition, *DIS3L2* is a candidate gene for Perlman syndrome, characterized by craniofacial abnormalities. Moreover, the frequency of the derived C allele is higher in EAS (47.8%) than in EUR (2.2%). The estimated allele age of this variant is about 7,875 years old (Table 1),

when EAS and EUR were already separated, suggesting that this variant arose in East Asian populations⁴⁹.

Interestingly, there are four independent variants associated with nose morphology in the *SOX9* locus in our EAS GWAS (Extended Data Fig. 7). Although *SOX9* is a well-known gene contributing to variation in nose shape, rs8068343 (NC_000017.10:g.69447706T>C) is a new variant that affects nose shape differences between populations^{7,9,10}. In contrast, the other three variants may affect nose shape within populations. Compared with the variants previously identified near *SOX9*, the reference (T) allele of rs8068343 has a lower frequency in EUR (0.04) than in EAS (0.46). Moreover, this variant had a higher F_{ST} (0.528) between EUR and EAS than any other variants, and the integrated haplotype score (iHS) of this variant was also higher in CHB (Han Chinese, 2.55)⁵⁰. These results indicate that this variant contributes to EUR–EAS nose shape difference.

Nose-associated variants are under positive selection. To assess whether the variation in facial morphology in EAS and EUR populations is due to natural selection or random drift, we conducted several selection analyses of the leading variants discovered here. The F_{ST} enrichment test showed that regions of the whole face and nose had a significantly higher F_{ST} than random variants after Bonferroni correction ($P_{\text{whole face}} = 8.22 \times 10^{-7}$, $P_{\text{nose}} = 1.00 \times 10^{-4}$; Supplementary Table 9 and Fig. 4a), indicating that facial morphology has been under natural selection in EAS and EUR, especially in the nose region⁵¹. XP-EHH enrichment analysis showed a consistent result ($P_{\text{whole face}} = 3.78 \times 10^{-3}$, $P_{\text{nose}} = 1.27 \times 10^{-3}$; Supplementary Table 9 and Fig. 4b).

Next, the mean PBS values for nose-associated loci were significantly higher than random variants in EUR ($P = 6.90 \times 10^{-4}$) but not in EAS ($P = 0.15$) (Supplementary Table 9 and Fig. 4c), indicating that nose shape may be under subtle local selection in Europeans rather than in East Asians. In addition, we found that mean PBS values for the nose-associated loci in a recent published EUR facial GWAS were significantly higher in EUR ($P_{\text{EUR}} = 9.46 \times 10^{-3}$) but not in EAS ($P_{\text{EAS}} = 0.464$; Fig. 4d)¹⁸. The results support that nose shape may be under local selection in EUR rather than in EAS populations.

Based on a study by He et al., we estimated and tested differences using the selection coefficients for nose variants with positive contribution to East Asian nose appearance (EAS-FA in the nose segment)⁵². The EUR population showed higher selection coefficients

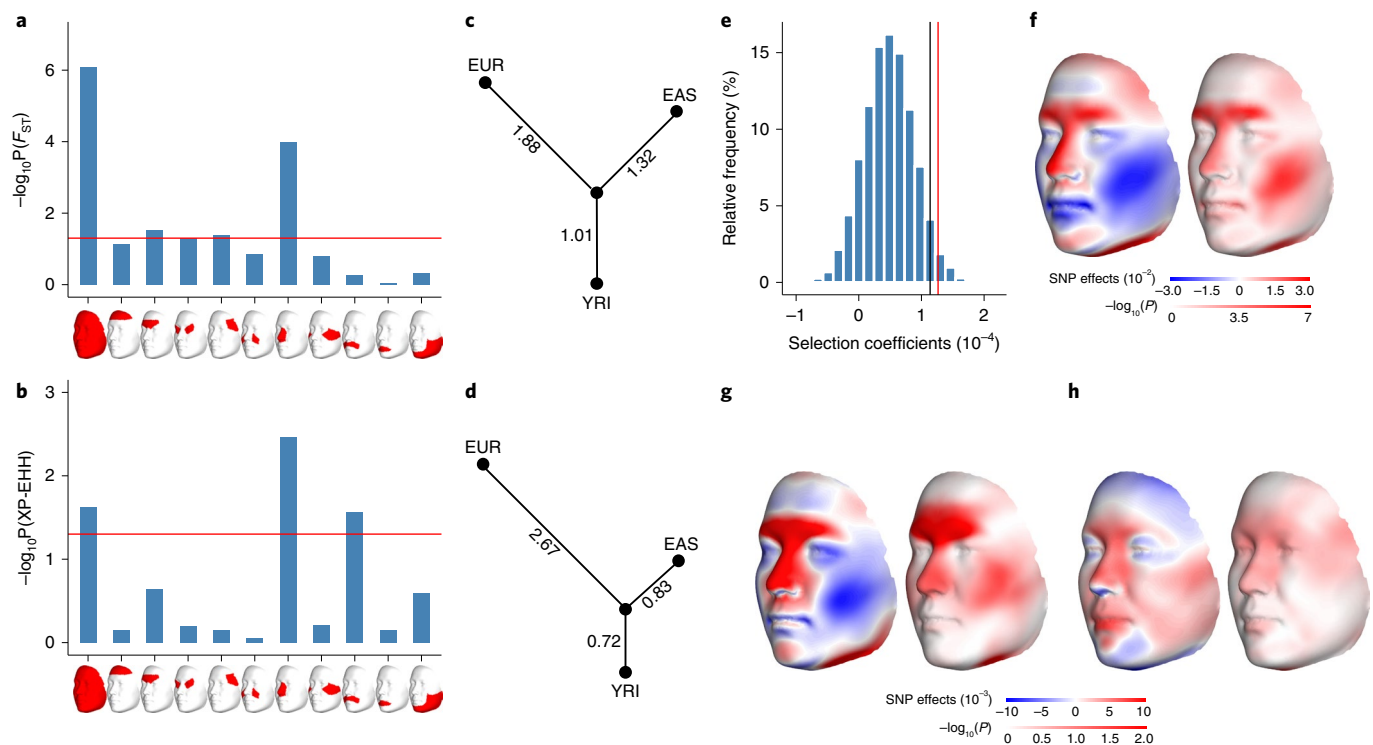


Fig. 4 | Natural selection analyses and enrichment test of the differentiation of facial-associated variants among EAS and EUR populations. **a, b**, P values ($-\log_{10}$ scale) of F_{ST} (**a**) and XP-EHH (**b**) for the whole face and ten anatomical regions. The red line is the P value threshold of 0.05. **c, d**, Observed mean PBS value for the leading variants among the 244 variants in this EAS study (**c**), and the 203 variants from study of White et al.¹⁸ (**d**) against the null distribution among EAS, EUR and YRI for the nose region. **e**, Selection coefficients for the nose region against the underlying null distribution (blue). The red line corresponds to the observed selection coefficients. The black line is 95% quantile of the null distribution. **f**, Differentiated accumulated genetic effects of the 244 leading variants (visualized using the local surface normal displacement) and P values ($-\log_{10}$ scale) of each quasi-landmark. **g, h**, Effects and P values ($-\log_{10}$ scale) of each quasi-landmark compared with random drift in the European (**g**) and East Asian (**h**) populations.

for nose EAS-FA increasing variants than the EAS population ($P = 4.88 \times 10^{-2}$; Supplementary Table 9 and Fig. 4e). Moreover, by comparing the mean genetic prediction of facial variation in EAS and EUR to the expected difference under random drift (Methods), the nose and glabella morphology in EUR was more protruding than in EAS and the divergence of the nose was greater than expected under the neutral model (Fig. 4f)³⁴. Furthermore, by comparing the PPS using the leading variants with the expected PPS under random drift in the EUR and EAS, we obtained the direction (and significance) of natural selection on facial morphology in each population³⁴. Similar to the population differences, the nose, glabella and zygoma were under natural selection in EUR (Fig. 4g). However, in EAS, the effects of natural selection were weaker (Fig. 4h). These results suggest that facial morphology in EUR may undergo local adaptation, producing a more protruded nose and glabella, and flatter zygoma.

On the basis of the above results, we speculate that facial features underlying EUR-EAS differences are potentially due to the adaptive selection that occurred in the EUR population, which makes European-ancestry populations have protruded and narrow noses, noticeably different from those of East Asian-ancestry populations.

Discussion

As a large-scale East Asian population facial GWAS using data-driven global-to-local phenotyping, our study broadens the knowledge of craniofacial genetics outside frequently investigated European-ancestry populations. Compared with previous facial GWASs, we identified 130 (out of 244) new variants associated with typical-range facial variation, which have similar biological functions as variants identified previously^{1–18}. A considerable number

of shared variants were identified independently in a EUR study and our EAS study using the same facial phenotyping approach. Among the 114 known variants, 96 were associated with consistent facial regions reported in previous facial GWASs. When we compared shared variants with the EUR study of White et al., 82 out of 89 were associated with the same facial regions, which indicates that different segmentation patterns could yield similar GWAS results (Supplementary Tables 3 and 6)¹⁸. These results suggest that the 244 variants identified in our study are reliable and that the genetic factors underlying facial variation might be universal across populations.

We further extended the concepts of polygenic scores (PGS) to polygenic shapes (PS) to verify whether the association found in EAS could be generalized to EUR³⁷. Both visual and statistical evidence supported this hypothesis on the whole face and major anatomical facial regions. However, the PPS of the mandible, forehead and lower mouth showed some difference from the corresponding EUR and EAS average shapes, due mainly to the insufficient number of significant variants associated with these facial regions. Besides smaller phenotypic variation in these regions in EAS, environmental factors contributing to facial variation may also impact the results. In addition, the Q_{ST} analysis suggests that mandible, lower mouth and forehead exhibit fewer signals of facial shape differences between EAS and EUR¹⁹. This could explain some reason for the inconsistency of the PPS for these facial regions. Of future interest is to calculate the PPS by combining all the variants identified in EAS and EUR studies. This might further improve the PPS in representing population facial shapes.

Our study also provides insights into the genetic basis of the facial shape differences between European-ancestry populations and East

Asian-ancestry populations. In addition to identifying 13 primary variants contributing to EUR–EAS facial differences, we provided a method to investigate the genetic factors associated with interpopulation facial variation. These 13 variants all had positive and larger effects on EUR–EAS facial differences, shaping the faces of East Asian populations to be more EAS-FA. Again, corresponding with the PPS results, due to the innate limitation of GWAS, our study may overlook rare or fixed variants that lead to more EAS-FA in the EAS population. By applying our method to EUR populations, additional variants affecting EAS-FA might be discovered. Moreover, for those rare or fixed variants with opposite alleles between EUR and EAS, a single population GWAS has limitations and an admixture population is instead needed.

Due to the large number of significant variants identified in the nose region, our evolutionary analysis further supports the hypothesis proposed by Zaidi et al. that human nose shape has evolved in response to selection pressures⁵³. Again, the PBS analyses showed that nose shape difference between EUR and EAS is due mainly to natural selection in European-ancestry populations rather than in East Asian-ancestry populations¹⁸.

In conclusion, this study presents a large EAS population GWAS on 3D facial shapes. Our study identified a large number of previously unknown variants associated with normal-range facial shape variation. Using newly introduced polygenic shapes, we successfully depicted perceptually recognizable population average faces, making our results more tangible, comprehensive and intuitive. We identified 13 variants contributing to more EAS-FA and revealed natural selection in shaping EUR–EAS nose shape difference. Our findings will greatly facilitate the understanding of human facial morphology across populations.

Online content

Any methods, additional references, Nature Research reporting summaries, source data, extended data, supplementary information, acknowledgements, peer review information; details of author contributions and competing interests; and statements of data and code availability are available at <https://doi.org/10.1038/s41588-022-01038-7>.

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References

- Liu, F. et al. A genome-wide association study identifies five loci influencing facial morphology in Europeans. *PLoS Genet.* **8**, e1002932 (2012).
- Paternoster, L. et al. Genome-wide association study of three-dimensional facial morphology identifies a variant in PAX3 associated with nasion position. *Am. J. Hum. Genet.* **90**, 478–485 (2012).
- Adhikari, K. et al. A genome-wide association study identifies multiple loci for variation in human ear morphology. *Nat. Commun.* **6**, 7500 (2015).
- Adhikari, K. et al. A genome-wide association scan implicates DCHS2, RUNX2, GLI3, PAX1 and EDAR in human facial variation. *Nat. Commun.* **7**, 11616 (2016).
- Cole, J. B. et al. Genomewide association study of African children identifies association of SCHIP1 and PDE8A with facial size and shape. *PLoS Genet.* **12**, e1006174 (2016).
- Pickrell, J. K. et al. Detection and interpretation of shared genetic influences on 42 human traits. *Nat. Genet.* **48**, 709–717 (2016).
- Shaffer, J. R. et al. Genome-wide association study reveals multiple loci influencing normal human facial morphology. *PLoS Genet.* **12**, e1006149 (2016).
- Lee, M. K. et al. Genome-wide association study of facial morphology reveals novel associations with FREM1 and PARK2. *PLoS One* **12**, e0176566 (2017).
- Cha, S. et al. Identification of five novel genetic loci related to facial morphology by genome-wide association studies. *BMC Genomics* **19**, 481 (2018).
- Claes, P. et al. Genome-wide mapping of global-to-local genetic effects on human facial shape. *Nat. Genet.* **50**, 414–423 (2018).
- Crouch, D. J. M. et al. Genetics of the human face: identification of large-effect single gene variants. *Proc. Natl Acad. Sci. USA* **115**, E676–E685 (2018).
- Qiao, L. et al. Genome-wide variants of Eurasian facial shape differentiation and a prospective model of DNA based face prediction. *J. Genet. Genomics* **45**, 419–432 (2018).
- Li, Y. et al. EDAR, LYPLAL1, PRDM16, PAX3, DKK1, TNFSF12, CACNA2D3, and SUPT3H gene variants influence facial morphology in a Eurasian population. *Hum. Genet.* **138**, 681–689 (2019).
- Wu, W. et al. Whole-exome sequencing identified four loci influencing craniofacial morphology in northern Han Chinese. *Hum. Genet.* **138**, 601–611 (2019).
- Xiong, Z. et al. Novel genetic loci affecting facial shape variation in humans. *eLife* **8**, e49898 (2019).
- Huang, Y. et al. A genome-wide association study of facial morphology identifies novel genetic loci in Han Chinese. *J. Genet. Genomics* **48**, 198–207 (2021).
- Bonfante, B. et al. A GWAS in Latin Americans identifies novel face shape loci, implicating VPS13B and a Denisovan introgressed region in facial variation. *Sci. Adv.* **7**, eabc6160 (2021).
- White, J. D. et al. Insights into the genetic architecture of the human face. *Nat. Genet.* **53**, 45–53 (2021).
- Guo, J. et al. Variation and signatures of selection on the human face. *J. Hum. Evol.* **75**, 143–152 (2014).
- Li, J. & Ji, L. Adjusting multiple testing in multilocus analyses using the eigenvalues of a correlation matrix. *Heredity (Edinb.)* **95**, 221–227 (2005).
- Kanai, M., Tanaka, T. & Okada, Y. Empirical estimation of genome-wide significance thresholds based on the 1000 Genomes Project data set. *J. Hum. Genet.* **61**, 861–866 (2016).
- Stouffer, S. A., Suchman, E. A., De Vinney, L. C., Star, S. A. & Williams, R. M. Jr. *The American Soldier: Adjustment During Army Life* (Wiley, 1965).
- Genomes Project, C. et al. A global reference for human genetic variation. *Nature* **526**, 68–74 (2015).
- McLean, C. Y. et al. GREAT improves functional interpretation of cis-regulatory regions. *Nat. Biotechnol.* **28**, 495–501 (2010).
- Watanabe, K., Taskesen, E., van Bochoven, A. & Posthuma, D. Functional mapping and annotation of genetic associations with FUMA. *Nat. Commun.* **8**, 1826 (2017).
- Bernstein, B. E. et al. The NIH Roadmap Epigenomics Mapping Consortium. *Nat. Biotechnol.* **28**, 1045–1048 (2010).
- Kundaje, A. et al. Integrative analysis of 111 reference human epigenomes. *Nature* **518**, 317–330 (2015).
- Schmidt, E. M. et al. GREGOR: evaluating global enrichment of trait-associated variants in epigenomic features using a systematic, data-driven approach. *Bioinformatics* **31**, 2601–2606 (2015).
- Wilderman, A., VanOudenhove, J., Kron, J., Noonan, J. P. & Cotney, J. High-resolution epigenomic atlas of human embryonic craniofacial development. *Cell Rep.* **23**, 1581–1597 (2018).
- Som, P. M. & Naidich, T. P. Illustrated review of the embryology and development of the facial region, part 2: late development of the fetal face and changes in the face from the newborn to adulthood. *AJNR Am. J. Neuroradiol.* **35**, 10–18 (2014).
- Staley, J. R. et al. PhenoScanner: a database of human genotype-phenotype associations. *Bioinformatics* **32**, 3207–3209 (2016).
- Zhou, Y. et al. Metascape provides a biologist-oriented resource for the analysis of systems-level datasets. *Nat. Commun.* **10**, 1523 (2019).
- Weir, B. S. & Cockerham, C. C. Estimating F-statistics for the analysis of population structure. *Evolution* **38**, 1358–1370 (1984).
- Robinson, M. R. et al. Population genetic differentiation of height and body mass index across Europe. *Nat. Genet.* **47**, 1357–1362 (2015).
- Sabeti, P. C. et al. Genome-wide detection and characterization of positive selection in human populations. *Nature* **449**, 913–918 (2007).
- Gautier, M., Klassmann, A. & Vitalis, R. rehh 2.0: a reimplementation of the R package rehh to detect positive selection from haplotype structure. *Mol. Ecol. Resour.* **17**, 78–90 (2017).
- Choi, S. W., Mak, T. S. & O'Reilly, P. F. Tutorial: a guide to performing polygenic risk score analyses. *Nat. Protoc.* **15**, 2759–2772 (2020).
- Yi, X. et al. Sequencing of 50 human exomes reveals adaptation to high altitude. *Science* **329**, 75–78 (2010).
- Singh, M. K. et al. The T-box transcription factor Tbx15 is required for skeletal development. *Mech. Dev.* **122**, 131–144 (2005).
- Lausch, E. et al. TBX15 mutations cause craniofacial dysmorphism, hypoplasia of scapula and pelvis, and short stature in Cousin syndrome. *Am. J. Hum. Genet.* **83**, 649–655 (2008).
- Ding, H. L., Clouthier, D. E. & Artinger, K. B. Redundant roles of PRDM family members in zebrafish craniofacial development. *Dev. Dyn.* **242**, 67–79 (2013).
- Kamberov, Y. G. et al. Modeling recent human evolution in mice by expression of a selected EDAR variant. *Cell* **152**, 691–702 (2013).
- Tan, J. et al. The adaptive variant EDARV370A is associated with straight hair in East Asians. *Hum. Genet.* **132**, 1187–1191 (2013).

44. Shaffer, J. R. et al. Multiethnic GWAS reveals polygenic architecture of earlobe attachment. *Am. J. Hum. Genet.* **101**, 913–924 (2017).
 45. Gaudet, P., Livstone, M. S., Lewis, S. E. & Thomas, P. D. Phylogenetic-based propagation of functional annotations within the Gene Ontology consortium. *Brief. Bioinform.* **12**, 449–462 (2011).
 46. Zhao, Y. et al. Isolated cleft palate in mice with a targeted mutation of the LIM homeobox gene *lhx8*. *Proc. Natl Acad. Sci. USA* **96**, 15002–15006 (1999).
 47. Haenig, C. et al. Interactome mapping provides a network of neurodegenerative disease proteins and uncovers widespread protein aggregation in affected brains. *Cell Rep.* **32**, 108050 (2020).
 48. Astuti, D. et al. Germline mutations in *DIS3L2* cause the Perlman syndrome of overgrowth and Wilms tumor susceptibility. *Nat. Genet.* **44**, 277–284 (2012).
 49. Albers, P. K. & McVean, G. Dating genomic variants and shared ancestry in population-scale sequencing data. *PLoS Biol.* **18**, e3000586 (2020).
 50. Voight, B. F., Kudaravalli, S., Wen, X. & Pritchard, J. K. A map of recent positive selection in the human genome. *PLoS Biol.* **4**, e72 (2006).
 51. Guo, J. et al. Global genetic differentiation of complex traits shaped by natural selection in humans. *Nat. Commun.* **9**, 1865 (2018).
 52. He, Y. et al. A probabilistic method for testing and estimating selection differences between populations. *Genome Res* **25**, 1903–1909 (2015).
 53. Zaidi, A. A. et al. Investigating the case of human nose shape and climate adaptation. *PLoS Genet.* **13**, e1006616 (2017).
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Methods

Ethics statement. All participants provided written informed consent, and all study protocols were approved by the institutional review boards of the pertinent research institutions. The National Survey of Physical Traits (NSPT) is a subproject of The National Science & Technology Basic Research Project approved by the Ethics Committee of Human Genetic Resources of School of Life Sciences, Fudan University, Shanghai (14117). The Northern Han Chinese (NHC) cohort was approved by the Ethics Committee of Human Genetic Resources at the Shanghai Institute of Life Sciences, Chinese Academy of Sciences (ER-SIBS-261410-A1801). The Taizhou Longitudinal Study (TZL) was approved by the Ethics Committee of Human Genetic Resources at the Shanghai Institute of Life Sciences, Chinese Academy of Sciences (ER-SIBS-261410). Written informed consent was granted for each participant before enrollment in the study. We confirm that our study is compliant with the Guidance of the Ministry of Science and Technology (MOST) for the Review and Approval of Human Genetic Resources.

Sample and recruitment details. The samples in this study were collected from three independent cohorts, the NSPT cohort ($n = 3,322$), the NHC cohort ($n = 4,767$) and the TZL cohort ($n = 2,881$). For the NSPT sample, individuals were recruited at three Chinese cities: Nanning, Guangxi province ($n = 1,326$); Taizhou, Jiangsu Province ($n = 986$); Zhengzhou, Henan province ($n = 1,010$). In the NHC cohort, participants were recruited in Tangshan, Hebei province. These two cohorts constituted the discovery dataset. The TZL cohort, where individuals were recruited in Taizhou, Jiangsu province, were used as the replication dataset. The characteristics of the datasets are shown in Supplementary Table 2.

Genotyping and imputation. Since we used two different genotyping platforms in the discovery and replication datasets (details in the Supplementary Note), we chose to impute the two datasets separately, then combine the imputed results.

For each dataset, standard data cleaning and quality assurance practices were performed based on the GRCh37 genome assembly. We performed phasing using SHAPEIT2 (v.2.17)³⁴ and imputation to the 1000GP Phase 3 reference panel using IMPUTE2 (v.2.3.2)³⁵. After postimputation quality control, 8,018,212 shared variants were obtained for analysis.

3D image acquisition, registration and quality control. The 3D images of all individuals in the three cohorts were captured and acquired using the 3dMDface (3dMD) camera system. When capturing, participants were asked to close their mouth, open their eyes and hold faces with a neutral expression.

The 3D surface images were registered using MeshMonk (v.0.0.6)³⁶ in MATLAB 2018a. This performed a homologous configuration of 7,906 spatially dense landmarks, allowing the 3D image data to be standardized. We performed generalized procrustes analysis (GPA) and symmetrization, then investigated every mapped image manually and identified outlier images. Any 3D facial images with poor quality were removed or reprocessed, with details available in the Supplementary Note.

As a result, 6,968 ($n = 4,089$ in the NHC cohort, $n = 2,879$ in the NSPT cohort) and 2,706 unrelated individuals with good quality 3D images in the discovery and replication dataset were used for further analysis.

Facial phenotyping. Like the approach of White et al., we performed a semisupervised facial segmentation based on the phenotypic correlation between facial landmarks using the discovery dataset¹⁸. To calculate the phenotypic correlations, we first corrected the symmetrized facial shapes for the covariates of age, age squared, sex, body mass index (BMI) and four SUGIBS components using a partial least-squares regression (PLSR, function `plsregress` from MATLAB 2018a) in both the discovery and replication cohort¹⁸. SUGIBS is a method for genetically robust genome-wide ancestry inference based on the spectral (S) decomposition of an unnormalized genomic (UG) relationship matrix generalized by an identity-by-state (IBS) similarity degree of the matrix of individuals, which was also used by White et al.^{18,57}

To study global and local effects on facial variation, we refined the data-driven facial segmentation method to avoid isolated point- and cluster-specific facial regions. We performed a hierarchical spectral clustering on a combined matrix, as $0.9 \times RV$ similarity matrix + $0.1 \times$ distance matrix, up to level five, resulting in a total of 63 facial segments (Supplementary Fig. 6 and Supplementary Note). In each segment, we performed principal component analysis (PCA) on the PLSR residuals of the discovery cohort and obtained principal component (PC) scores as the phenotypic scores for the discovery cohort. In the replication cohort, we projected the PLSR residuals onto the PCA space built in the discovery cohort to obtain PC scores as phenotypic scores. We described the methods in detail in the Supplementary Note.

Multivariate genome-wide association meta-analyses. The association analysis is similar to that in White et al.¹⁸. For all three phases (discovery, replication and meta-analysis), the genotypes were coded as the number of major alleles present (0, 1 or 2). In the discovery phase, in each of the 63 facial segments, we used canonical correlation analysis (CCA) to define the linear combination of the facial segments PCs that are most correlated with each variant, which represent the

phenotypic effect in shape space. When one of the two sets of variables has only one variable, CCA reduces to multiple regression⁵⁸. The resulting vector ω_i is the effect size vector of this variant in the shape PCA space (Supplementary Note). The correlation can be tested for significance based on Rao's exact F -test (one-sided, right tail)⁵⁹. For each variant, we obtained a direction ω_i in the shape PCA space most correlated with the genotype of that variant and a P value representing the strength of correlation in the discovery phase. In the replication phase, we first projected the phenotypic scores onto the CCA direction and calculated Pearson's correlation between the projected scores and the genotypes in the replication cohort. To test the correlation's significance, we used Student's t -test where the t -statistic is defined as $t = \frac{\sqrt{r^2(1-r^2)}}{\sqrt{n-2}}$. We performed a one-sided right-tail test for each variant to ensure that the effective direction of the variant in the two datasets is the same. Next, the P values obtained in the discovery and replication phase were combined in a meta-analysis using Stouffer's method weighted by the sample sizes²². We used the corresponding implementations of these methods in the SNPLIB package (available at <https://github.com/jiarui-li/SNPLIB>) to accelerate the analyses.

Conditional analysis and GWAS peak selection. For every variant, the meta-analysis described above yielded 63 P values representing 63 facial segments. In the conditional analysis and peak selection, we selected the lowest P value for each variant. For the initial selection, we selected the variants with a P value below the genome-wide threshold ($P = 5 \times 10^{-8}$) and calculated the pairwise r^2 between these variants. In each chromosome, we grouped the selected variants consecutively in such a way that the r^2 between every two neighbor-selected variants in the group is greater than 0.05, which resulted in 230 groups. Then, we selected the variant with the lowest P value as the conditional variant for each group, and performed association tests of the remaining variants on the condition of the conditional variant. The variant with the most significant P value that was still lower than the genome-wide threshold was selected as a conditional variant. We repeated these two steps until no variant remained significant. Finally, we obtained 244 leading single nucleotide polymorphisms (SNPs) from all groups. We considered a ± 500 kb genomic region of each leading variant as a genomic locus. If nearby genomic regions overlapped, we merged them into one genomic locus.

Permutation test of study-wide P value threshold. To determine the study-wide Bonferroni P value threshold, we calculated the number of independent tests by both the eigenvalues of the correlation matrix of the segments and the permutation analysis scheme used in the study of White et al.^{18,20,21}. The numbers of independent tests obtained from the eigenvalues of the correlation matrix and the permutation analysis are 50 and 51.41, respectively. Here, we used the more stringent threshold $5 \times 10^{-8}/51.41 = 9.8 \times 10^{-10}$. The details can be found in the Supplementary Note.

Heritability of Facial Segments. In each facial segment, we first calculated the genome-wide heritability of each retained shape PCs using the discovery cohort and calculated the facial segment's heritability as the mean heritability of the PCs weighted by the eigenvalue (variance) of each PC. To calculate the genomic relationship matrix (GRM) for heritability estimation, we first removed SNPs based on high levels of pairwise LD by PLINK v.1.9 with window size of 50, step size of 5 bases and $r^2 > 0.1$, remaining 266,241 SNPs. The heritability of all facial segment PCs was estimated by GCTA^{60,61}.

Gene mapping and functional annotation. Candidate genes of the genome-wide-significant leading SNPs were first identified using the NCBI, HaploReg v.4.1, UCSC genome browser and Ensemble genome browser^{62–65}. We also used three gene-mapping criteria implemented in Functional Mapping and Annotation (FUMA, v.1.3.6) to identify the most likely candidate gene per leading variant²⁵. First, we mapped variants to genes based on physical distance (within a 10,000 basepair window) from the known protein-coding genes in the human reference assembly. Second, we included the genes which have a significant cis-expression quantitative trait locus (eQTL; 1 Mb distance to the leading variant) association with the leading variants, using ten tissue types from the GTEx v.8 database^{66–68}. We used an FDR of 0.05 to define significant eQTL associations. Finally, we also identified candidate genes for each leading variant if there is chromatin interaction. To further prioritize candidate genes, we limited interaction-mapped genes to those that interact with a predicted enhancer region identified in any of the 111 tissues or cell types from the Roadmap Epigenomics Mapping Consortium (ROADMAP) and/or a gene promoter region (from 250 bp upstream to 500 bp downstream of the transcription start site (TSS) and predicted by the ROADMAP to be a promoter region)²⁷. We expected that the resulting candidate genes are more likely to have a plausible biological function. We used an FDR of 1×10^{-6} to define significant interactions. To further narrow down the candidate genes, we investigated whether any gene in the window was previously associated with craniofacial development or morphology through normal-range facial association studies, genetic disorders with facial dysmorphology phenotypes or animal models. To investigate the potential biological process of the candidate genes, FUMA (v.1.3.6) and GREAT (v.4.0.4) were performed using preset parameters^{24,25}.

Chromatin state association in embryonic craniofacial tissue. We used GREGOR (v.1.4.0) to evaluate global enrichment of trait-associated variants in different chromatin states²⁸. This method tests for an increase in the number of facial-associated index variants, or their LD proxies ($r^2 > 0.8$), overlapping with the regulatory feature more often than expected by chance by comparing with permuted control sets (random control variants are selected across the genome that match the index variant for several variants in LD, minor allele frequency and distance to nearest intron). The reference epigenomes of 127 human tissues and cell types were obtained from the NIH Roadmap Epigenomics Mapping Consortium²⁶. The human embryonic craniofacial chromHMM states were obtained from each Carnegie stage by Wilderman et al.²⁹.

Gene expression enrichment analysis. We selected a set of transcriptome datasets from critical periods of mouse face formation that enable gene expression to be analyzed with respect to time, prominence and tissue layer (GSE62214). We evaluate the expression level (fold change) of the candidate genes compared with a set of control groups where the genes were selected randomly from the genome. Then we regressed the fold change of gene expression on time, prominence and tissue layer to test their associations.

EAS and EUR average facial shapes. We recruited 89 individuals with self-reported European ancestry (32 females and 57 males) aged between 16 and 57 years old in Shanghai²⁵. They were required to have complete European ancestry over the last three generations. Their 3D facial images were captured using the same protocol as used in the Chinese cohort. In each segment, we aligned these EUR facial shapes to the corresponding sample full Procrustes mean shape. We then calculated the male and female average facial shapes separately and used the average facial shapes of the two average shapes as the EUR average facial shapes. To calculate the EAS average facial shapes, we selected five individuals in the Han Chinese cohort with matched age and gender to the individuals in the EUR cohort. We finally selected 445 (5 × 89) individuals to calculate the EAS average facial shapes in the same manner as in the EUR cohort.

Polygenic shape analysis. One can calculate the effect size vector β_i in the original shape space by:

$$\beta_i = V_k \Sigma_k \omega_i \quad (1)$$

where Σ_k is a diagonal matrix of the largest k singular values and the column vectors of V_k are the corresponding k right singular vectors obtained from the PCA (Supplementary Note).

Thus, the PS of an individual could be calculated as:

$$PS = \sum_i^n \beta_i g_i \quad (2)$$

where g_i is the genotype value of variant i (ref.³⁷). Subsequently, we calculate the PPS by:

$$PPS = 2 \sum_i^n \beta_i a_i \quad (3)$$

where a_i is the effect allele frequency of variant i , and two times a_i is the average number of effect alleles in a given population.

Next, we used the PPS difference between EUR and EAS ($PPS_{EUR} - PPS_{EAS}$) calculated by leading variants compared with random variants to evaluate whether leading variants could effectively fit the EUR–EAS shape difference. We calculated the PPS-derived shapes as follows:

$$F_{EAS}^d = F_{AVG} - \frac{PPS_{EUR} - PPS_{EAS}}{2} \quad (4)$$

$$F_{EUR}^d = F_{AVG} + \frac{PPS_{EUR} - PPS_{EAS}}{2} \quad (5)$$

where F_{EAS}^d and F_{EUR}^d are the corresponding PPS-derived shapes, F_{AVG} is the average facial shape of the population average shapes of EUR and EAS (that is, a population neutral average face).

We performed 1,000 simulations to calculate the random $PPS_{EUR} - PPS_{EAS}$. In each simulation, random variants with the same effect allele frequencies in EAS were chosen to calculate random $PPS_{EUR} - PPS_{EAS}$. Subsequently, the cosine similarity and the Euclidean distances between the PPS and the corresponding average face were used as measures of shape similarity. P values of each approach were then calculated using the null distribution established by these 1,000 simulations.

Contribution of variants to EAS–FA. We used the projected (vector) length to quantify the contribution of a variant to the EUR–EAS face difference:

$$l_i = \frac{2(a_i^{EUR} - a_i^{EAS})\beta_i(\bar{F}_{EUR} - \bar{F}_{EAS})}{|\bar{F}_{EUR} - \bar{F}_{EAS}|} \quad (6)$$

If a variant has a positive sign of projected length, we regard this variant to be linked to EAS individuals having more EAS features. In contrast, a variant with a negative sign is linked with EAS individuals having more EUR characteristic features.

Genetic information from 1000 Genome Project. We used the individuals' genetic information from 1000GP Phase 3 for related analysis including calculating MAF, PPS analysis, F_{ST} and PBS calculation. EAS including Han Chinese in Beijing (CHB, $n = 103$) and Southern Han Chinese (CHS, $n = 105$); EUR including Toscani (TSI, $n = 107$), British (GBR, $n = 91$), Iberian (IBS, $n = 107$) and Utah residents (CEU, $n = 99$), without Finnish (FIN); and AFR including Yoruba (YRI, $n = 103$) were used for analysis.

Calculation of natural selection signatures. We calculated genome-wide natural selection signatures based on XP-EHH using REHH2 (v.3.2.0)³⁶. The genome-wide XP-EHH z -scores were standardized through normalization in each derived allele frequency bin (bin widths = 0.01). We estimated two-tailed P values of the variant according to the normalized z -scores. We calculated the F_{ST} and PBS for different sets of a population, by using the observed allele frequencies estimated from the 1000GP Phase 3 (refs.^{33,38}). On the basis of a previous study, we measured selective pressures by (genic) selection coefficients, the details of those calculations are described in He et al.³².

Phenome-wide selection signature analysis. Similar to the approach used in Guo et al., we compared the mean F_{ST} /PBS value of the leading variants with that of the control variants with MAF and LD score matched⁵¹. First, we divided all the variants (1000GP) into 20 MAF bins from 0 to 0.5 with an increment of 0.025 (excluding the SNPs with MAF < 0.01). Each of the MAF bins was further grouped into 20 bins according to the 20 quantiles of LD score distribution. The MAF and LD score values were computed from the EAS or EUR samples in the 1000GP described above. Second, we allocated the leading variants to the MAF and LD stratified bins, randomly sampled a matched number of 'control' variants from each bin, computed a mean F_{ST} /PBS value for the control variants sampled from all bins, and repeated this process 10,000 times to generate a distribution of mean F_{ST} /PBS under drift. Third, a P value was computed from a two-tailed test by comparing the observed mean F_{ST} /PBS value for the leading variants against the null distribution quantified by the control variants, assuming normality of the null distribution. Regarding enrichment analysis of the selection signatures by XP-EHH, we obtained the sum of the squared values of the normalized XP-EHH z -scores of the variants (or the proxy variants in LD when available; $r^2 > 0.6$ in the CHB or CEU data from 1000 Genome Project), which was compared with the χ^2 distribution with the degree of freedom equal to the number of the variants.

Direction of genetic differentiation. The analysis presented here uses a method similar to that introduced in Robinson et al. to quantify the population genetic differentiation of a complex trait³⁴. The coefficients of the leading variants were randomized across variants 10,000 times, and 10,000 genetic predictors were created in the EAS or EUR samples from the 1000GP described above. By keeping the effect sizes consistent but attributing these effects across variants at random, the genetic predictors generated reflect the action of genetic drift.

Reporting Summary. Further information on research design is available in the Nature Research Reporting Summary linked to this article.

Data availability

The Meta-analysis GWAS summary statistics are available on the National Omics Data Encyclopedia. NODE: OEP002283. The participants making up the NSPT, NHC and TZL datasets were not collected with broad data sharing consent. Given the highly identifiable nature of both facial and genomic information and unresolved issues regarding risk to participants, we opted for a more conservative approach to participant recruitment. Broad data sharing of the raw data from these collections would thus be in legal and ethical violation of the informed consent obtained from the participants. This restriction is not because of any personal or commercial interests. Additional details can be requested from L.J. for the NSPT dataset, and S. Wang for the NHC and TZL datasets. Data usage shall be in full compliance with the Regulations on Management of Human Genetic Resources in China. Publicly available data used were: the 1000GP Phase 3 data (<https://www.internationalgenome.org/category/phase-3/>)²³, The Roadmap Epigenomics Project (<http://www.roadmapepigenomics.org/>)²⁶, NCBI dbSNP database (<http://www.ncbi.nlm.nih.gov/SNP/>)⁶², UCSC genome browser (<http://genome.ucsc.edu/>)⁶³, HaploReg v.4.1 (<https://pubs.broadinstitute.org/mammals/haploreg/haploreg.php>)⁶⁴, Ensemble genome browser (http://asia.ensembl.org/Homo_sapiens/Info/Index)⁶⁵, GTEx v.8 (<https://gtexportal.org/home/>)^{66–68}, Human genome dating (<https://human.genome.dating/>)⁴⁹ and the transcriptome resource from separated ectoderm and mesenchyme of the developing mouse face (GSE62214).

Code availability

The statistical analyses in this work were based on functions of the statistical toolbox in MeshMonk (<https://github.com/TheWebMonks/meshmonk>, v.0.0.6)³⁶,

MATLAB 2018a, R (v.3.6.1), ggplot2 (v.3.1.0), Python (v.3.5.0), PLINK v.1.9, SHAPEIT2 (v.2.17), IMPUTE2 (v.2.3.2), SNPLIB (<https://github.com/jiarui-li/SNPLIB>), GCTA-GREML, FUMA (v.1.3.6), GREAT (v.4.0.4), GREGOR (v.1.4.0), Metascape (<https://metascape.org>), LocusZoom (<https://genome.sph.umich.edu/wiki/LocusZoom>) and REHH2 (v.3.2.0) as mentioned throughout the Methods.

References

54. Delaneau, O., Marchini, J. & Zagury, J. F. A linear complexity phasing method for thousands of genomes. *Nat. Methods* **9**, 179–81 (2011).
55. Howie, B. N., Donnelly, P. & Marchini, J. A flexible and accurate genotype imputation method for the next generation of genome-wide association studies. *PLoS Genet.* **5**, e1000529 (2009).
56. White, J. D. et al. MeshMonk: Open-source large-scale intensive 3D phenotyping. *Sci. Rep.* **9**, 6085 (2019).
57. Li, J. et al. Robust genome-wide ancestry inference for heterogeneous datasets: illustrated using the 1,000 genome project with 3D facial images. *Sci. Rep.* **10**, 11850 (2020).
58. Rencher, A. C. & Christensen, W. F. *Methods of Multivariate Analysis*. (John Wiley & Sons, Inc., 2012). <https://doi.org/10.1002/9781118391686>
59. Olson, C. L. On choosing a test statistic in multivariate analysis of variance. *Psychol. Bull.* **83**, 579–586 (1976).
60. Yang, J. et al. Common SNPs explain a large proportion of the heritability for human height. *Nat. Genet.* **42**, 565–569 (2010).
61. Yang, J., Lee, S. H., Goddard, M. E. & Visscher, P. M. GCTA: a tool for genome-wide complex trait analysis. *Am. J. Hum. Genet.* **88**, 76–82 (2011).
62. Sherry, S. T. et al. dbSNP: the NCBI database of genetic variation. *Nucleic Acids Res.* **29**, 308–311 (2001).
63. Navarro Gonzalez, J. et al. The UCSC Genome Browser database: 2021 update. *Nucleic Acids Res.* **49**, D1046–D1057 (2021).
64. Ward, L. D. & Kellis, M. HaploReg: a resource for exploring chromatin states, conservation, and regulatory motif alterations within sets of genetically linked variants. *Nucleic Acids Res.* **40**, D930–4 (2012).
65. Aken, B. L. The Ensembl gene annotation system. *Database (Oxford)* **2016**, baw093 (2016).
66. Lonsdale, J. et al. The Genotype-Tissue Expression (GTEx) project. *Nat. Genet.* **45**, 580–585 (2013).
67. Carithers, L. J. & Moore, H. M. The Genotype-Tissue Expression (GTEx) Project. *Biopreserv. Biobank.* **13**, 307–308 (2015).
68. Keen, J. C. & Moore, H. M. The Genotype-Tissue Expression (GTEx) project: linking clinical data with molecular analysis to advance personalized medicine. *J. Pers. Med.* **5**, 22–29 (2015).

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Author contributions

S. Wang, L.J., J.L. and M.Z. conceptualized the study (ideas; formulation or evolution of overarching research goals and aims). M.Z., S. Wu, S.D., W.Q. and J.L. carried out data curation (management activities to annotate, scrub data and maintain research data for initial use and later re-use). M.Z., S. Wu, S.D., W.Q., J.L. and J.C. carried out the formal analysis (application of statistical, mathematical, computational or other formal techniques to analyze or synthesize study data). M.Z., S. Wu, S.D., W.Q. and J.L. did the visualization (preparation, creation and/or presentation of the published work, specifically visualization/data presentation). K.T., L.Q., Y.Y. and J.T. collected the 3D facial scans of the TZL cohort. J.W., Z.Y., J.T., K.T. and L.Q. collected the 3D facial scans of the NSPT cohort. S. Wu, Y.L. and Q.P. contributed to generating the SNP array data. S. Wu, S.D. and J.L. registered the 3D facial scans of the Northern Han Chinese cohort and conducted the PCA of discovery cohort. N.N. and A.R.-L. performed the analysis of MPRS22/Mprs22 and human/mouse craniofacial shape. M.Z., S. Wu, S.D., W.Q. and J.L. wrote the original draft. S. Wang, L.J., P.C., J.L., M.Z., S. Wu, S.D. and W.Q. reviewed and edited the final manuscript. All authors participated in preparing the manuscript by reading and commenting on drafts before submission.

Competing interests

The authors declare no competing interests.

Additional information

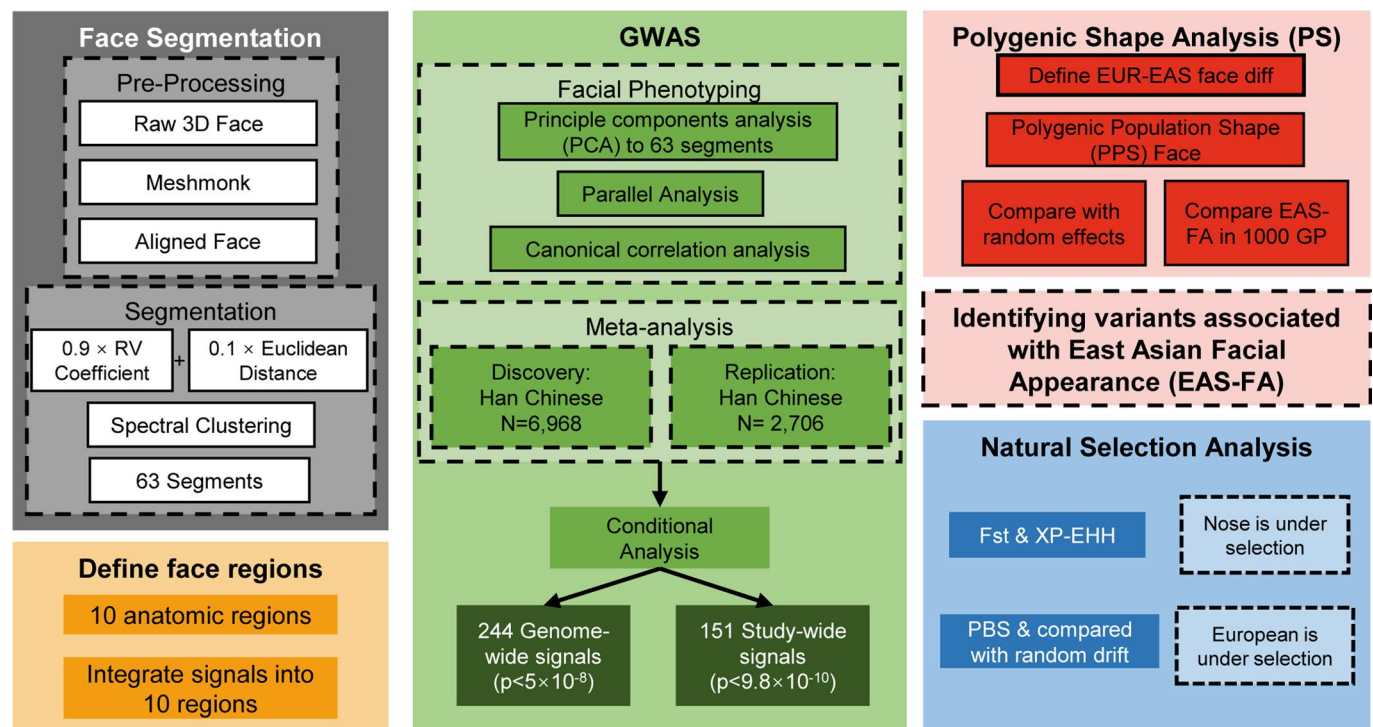
Extended data is available for this paper at <https://doi.org/10.1038/s41588-022-01038-7>.

Supplementary information The online version contains supplementary material available at <https://doi.org/10.1038/s41588-022-01038-7>.

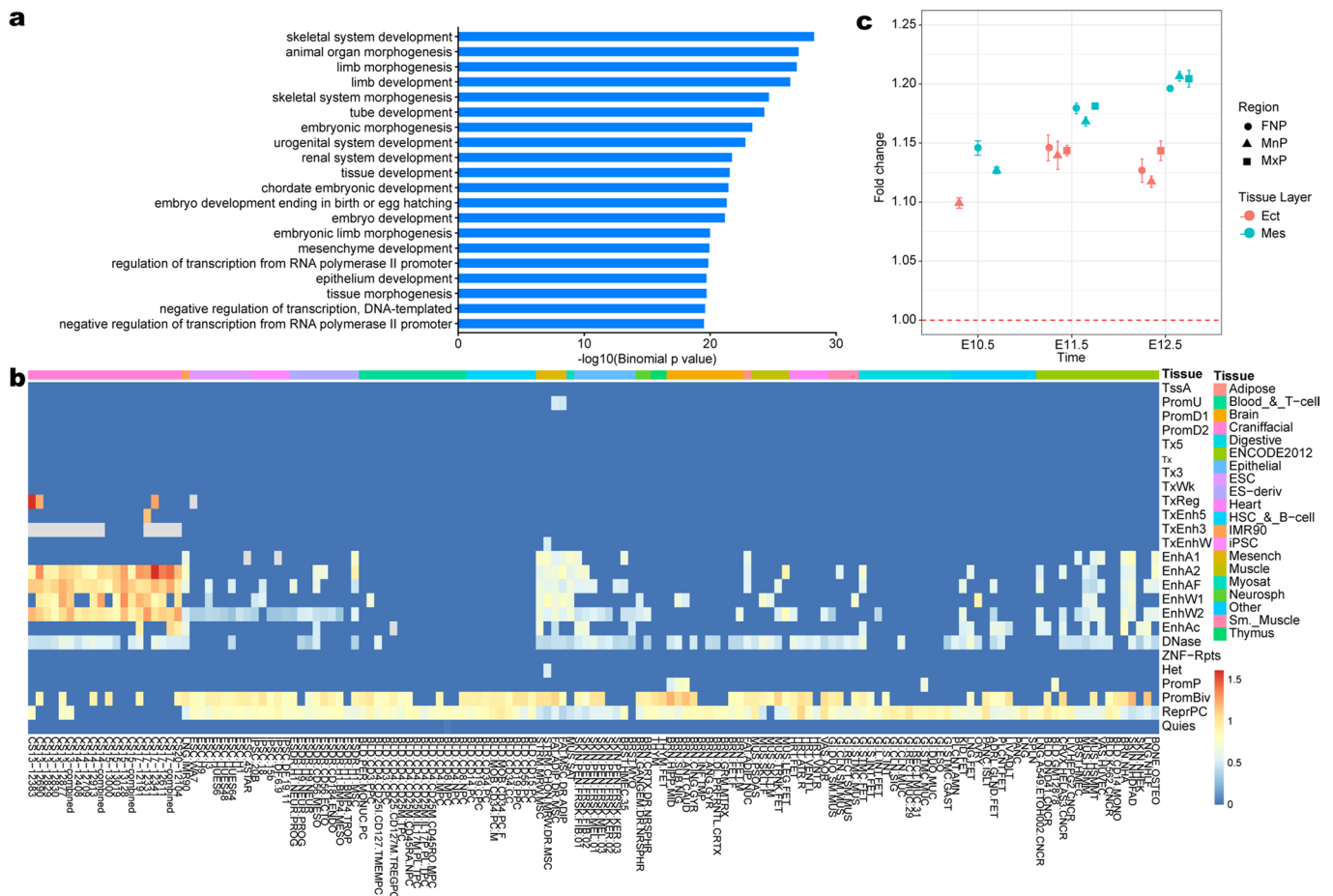
Correspondence and requests for materials should be addressed to Li Jin, Jiarui Li or Sijia Wang.

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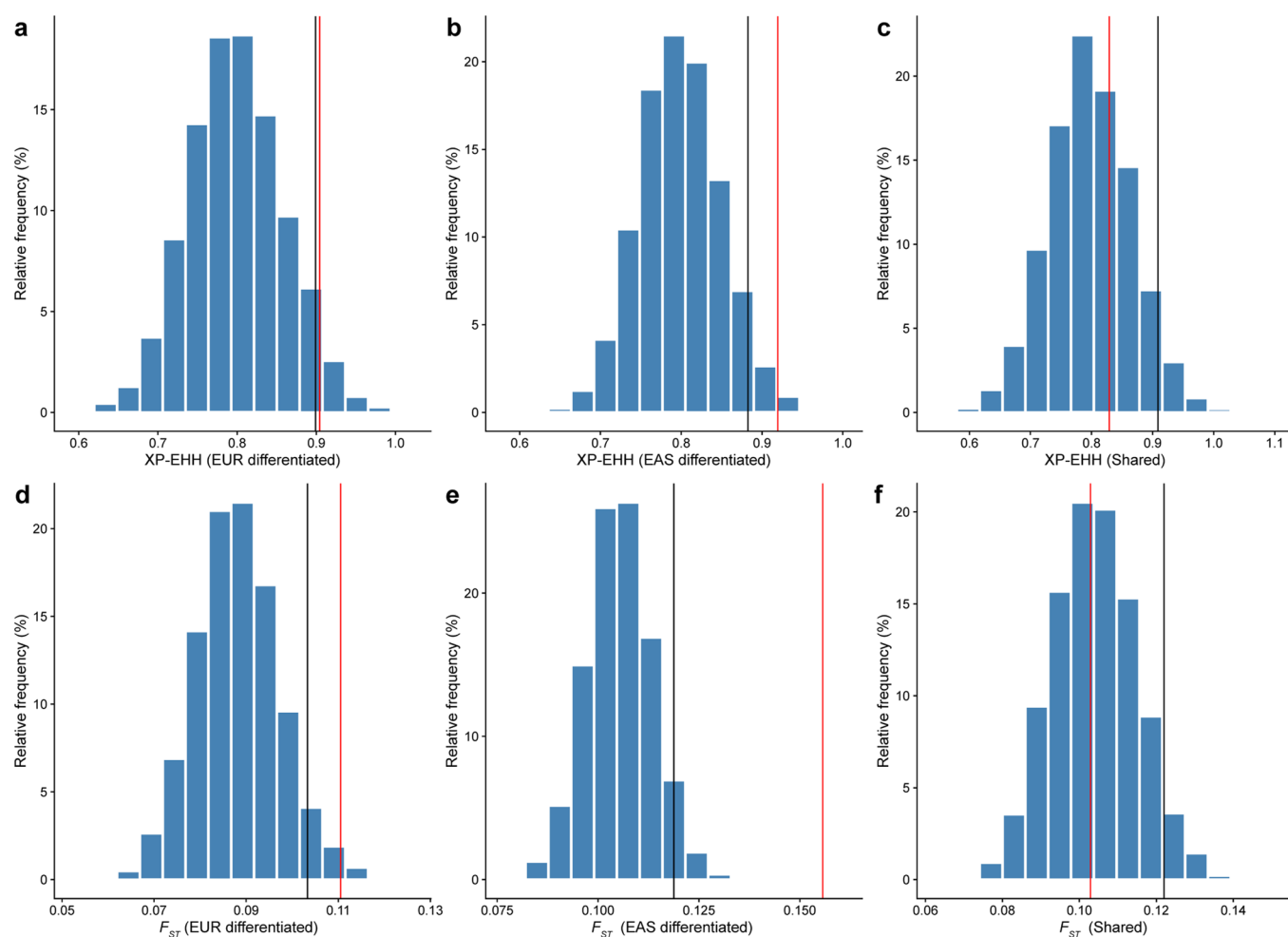
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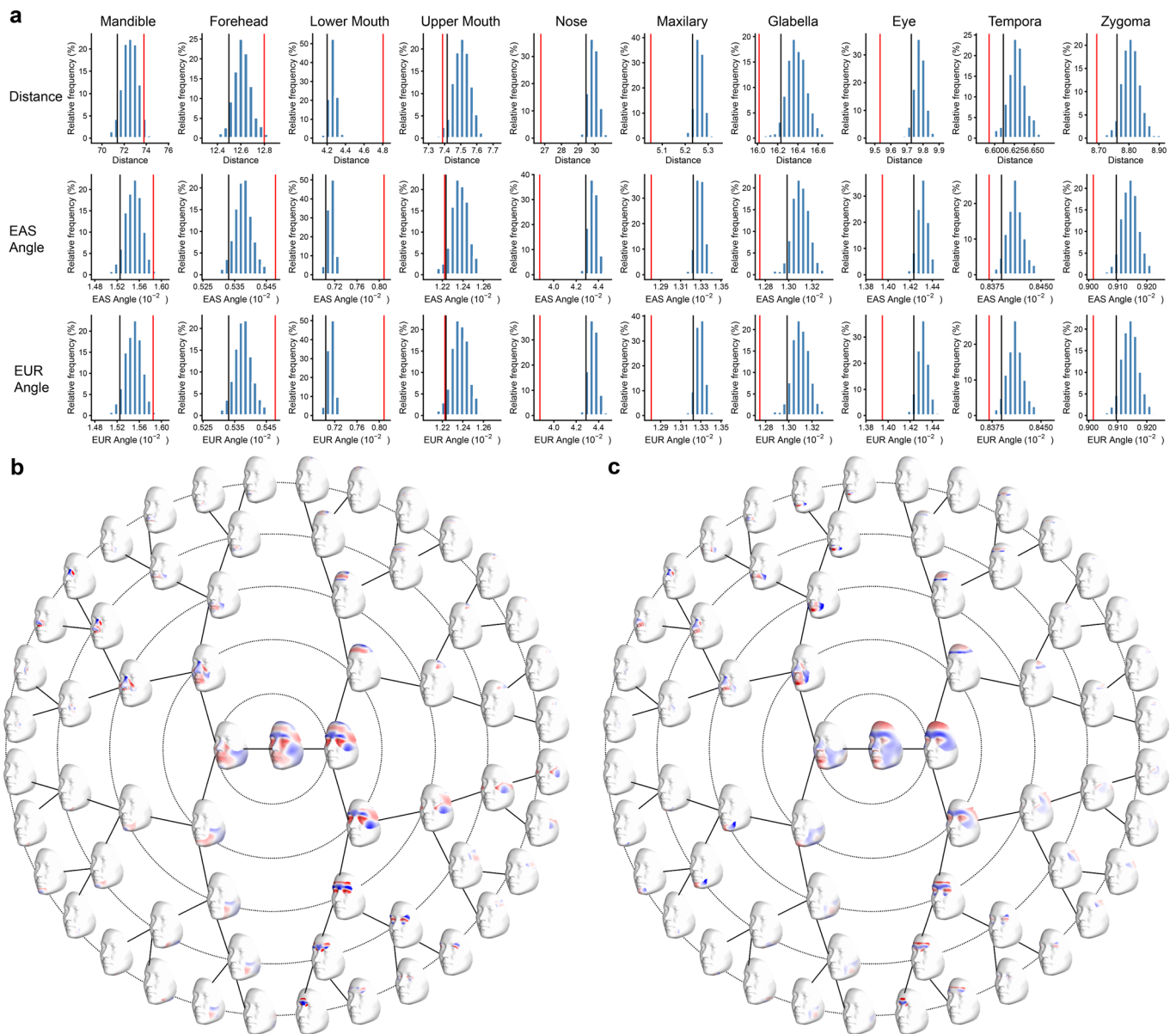
Extended Data Fig. 1 | Study design. We first start with a face segmentation procedure to get 63 face segments from which we defined 10 anatomical face regions. Then by using a CCA based GWAS, we identified 244 variants with a P value lower than 5×10^{-8} , in which 151 are also lower than 9.8×10^{-10} . To investigate what affects the similarity of an EAS face, we used polygenic population shape (PPS) analyses to fit EUR and EAS faces and identified 13 variants mainly contributing to EUR-EAS facial differences. To investigate selection on facial variation, we used F_{ST} and XP-EHH to find which parts of the face are under selection. These results, we further compared with random drift and random PPS to find out, which from the two populations, EUR or EAS, experienced selection.



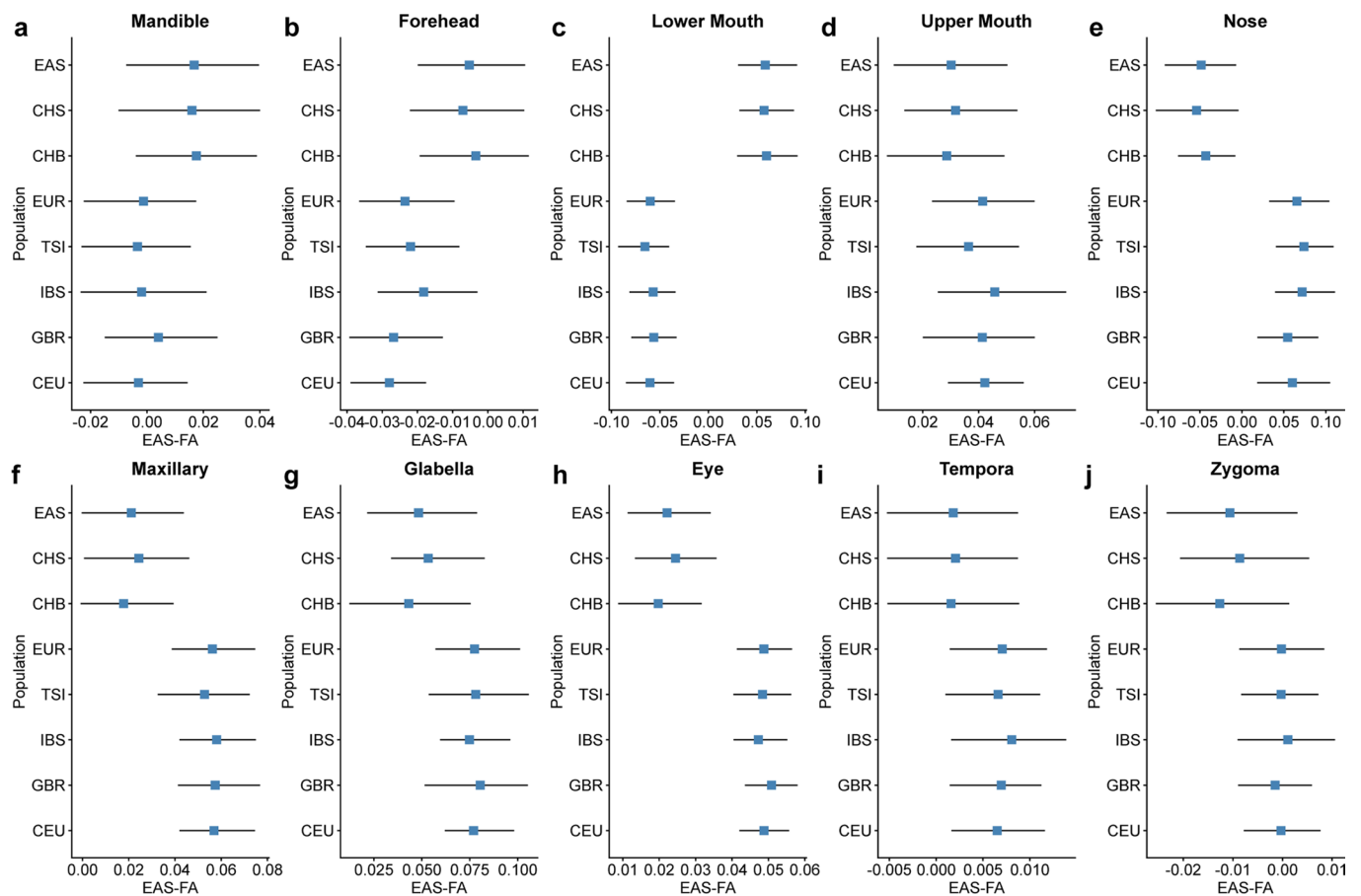
Extended Data Fig. 2 | Enrichment analysis of leading variants. (a) Gene Ontology enrichment for genes annotated from leading variants by GREAT²⁴. (b) Heatmap indicating the global enrichment of trait-associated variants in different chromatin state (y axis) and in different tissue (x axis). The fold change was calculated by GREGOR²⁸. The embryonic craniofacial tissue was previously published by epigenomic atlas, while the other was previously published by Roadmap Epigenome²⁷. The description of the 25-state chromatin model can be found at: https://egg2.wustl.edu/roadmap/web_portal/imputed.html#chr_imp. (c) Expression levels of the candidate genes in craniofacial tissues. Each point ($n=3$ biologically independent replicates for each condition) represents an estimated fold change compared to control genes at different times (E10.5, E11.5, E12.5), in different prominences (Frontonasal, FNP: circle; Maxillary, MxP: square; Mandibular, MnP: triangle), and tissue layer (Ectoderm, Ect: red; Mesenchyme, Mes: blue). Data are presented as mean values \pm 95% confidence intervals ($1.96 \times \text{SEM}$).



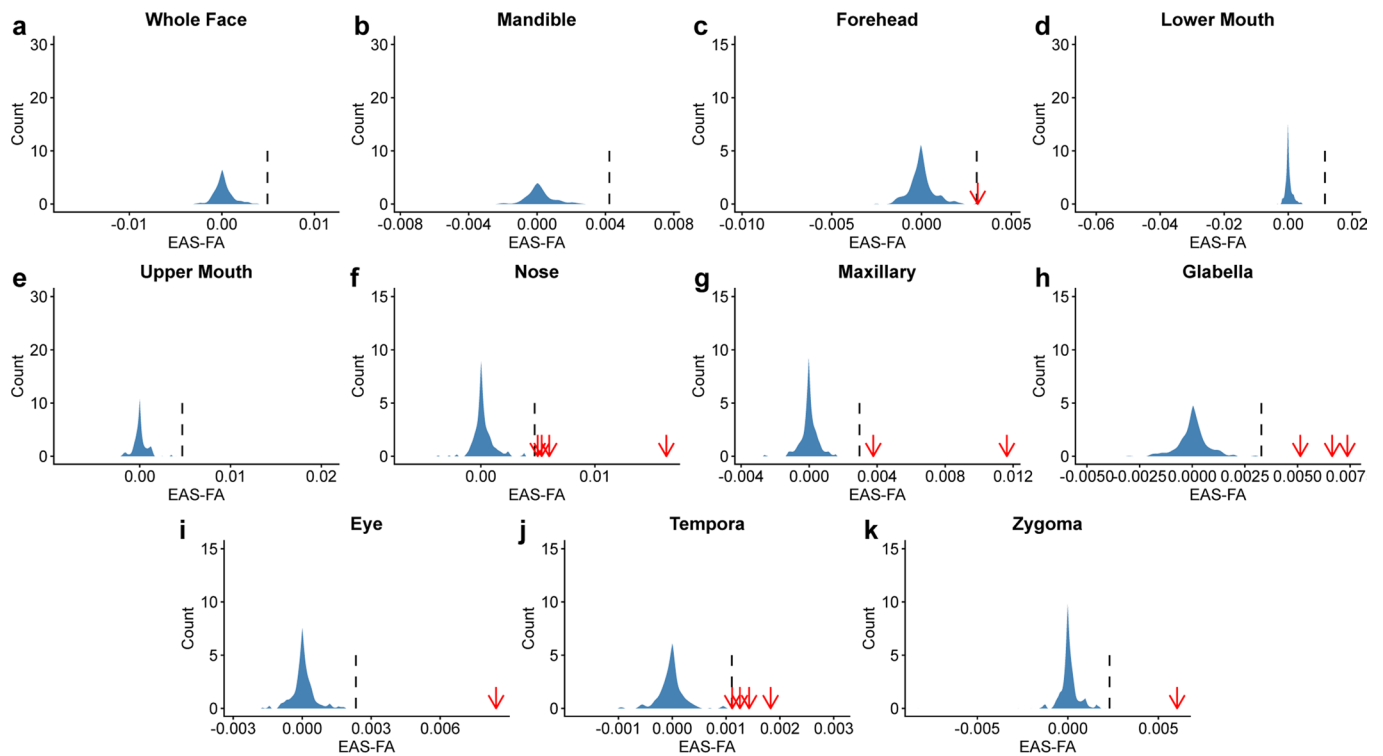
Extended Data Fig. 3 | XP-EHH and F_{ST} enrichment analysis for shared and differentiated variants. XP-EHH and F_{ST} enrichment analysis for (a, d) EUR differentiated variants, (b, e) EAS differentiated variants, and (c, f) shared variants in EAS study. The blue color is the null distribution. The red line is the mean XP-EHH or F_{ST} score of shared or differentiated variants. The black line is the 95% quantile of the null distribution.



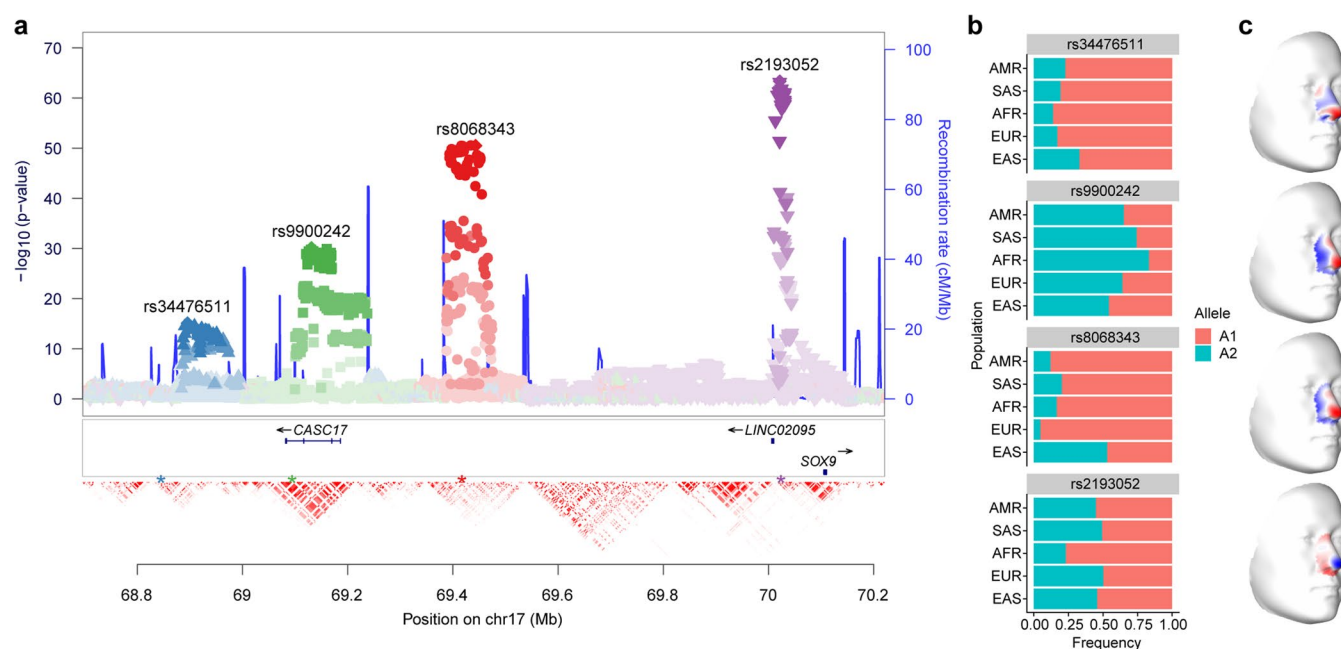
Extended Data Fig. 4 | Validation of PPS in 10 anatomical segments. (a) The null distribution (blue) of Euclidean distance, cosine similarity with EUR average face and EAS average face using 1,000 simulations derived from random variants on the 10 anatomical regions, red line infers the statistics of the leading variants associated with corresponding regions; black line infers 95% quantile of distribution from the random variants with corresponding regions; (b) The genetic effects of rs12632544 and (c) rs12473319 weighted by their effect allele number difference of EUR and EAS (visualized using the local surface normal displacement).



Extended Data Fig. 5 | The EAS-FA of polygenic shapes in 10 anatomical regions for EAS and EUR individuals in 1000GP. The EAS-FA of polygenic shapes in a) mandible, b) forehead, c) lower mouth, d) upper mouth, e) nose, f) maxillary, g) glabella, h) eye, i) tempora, and j) zygoma for EAS and EUR individuals in 1000GP. The squares represent the mean EAS-FA score in 10 anatomical regions and the horizontal lines represent 1st and 3rd quantile.



Extended Data Fig. 6 | EAS-FA of the 244 leading variants on the EUR-EAS difference. The distributions (blue) of EAS-FA derived from 244 leading variants associated with a) whole face and b) - k) 10 anatomical segments. The black dotted line is the EAS-FA threshold of each region (mean + 3xSD). The red arrow is the variant over threshold.



Extended Data Fig. 7 | Multi peak in 17q24.3 region. (a) Association variants in the SOX9 locus and genomic environment surrounding SOX9 across a 2-Mb window. Four independent variants, represented by (1) rs34476511 (blue), (2) rs9900242 (green), (3) rs8068343 (red), and (4) rs2193052 (purple) are observed; (b) Allele frequency in AMR, SAS, AFR, EUR and EAS population of the four variants from 1000GP; (c) The effects of the four variants in the nose region.

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Software and code

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Data collection	Three-dimensional images composed of surface and texture maps were taken using the 3dMD Face camera system (3dMD, Atlanta, GA). Genotyping was performed using the Illumina Infinium Global Screening Array or illumine HumanOmniZhongHua-8 array.
Data analysis	The statistical analyses in this work were based on functions of the statistical toolbox in MeshMonk (https://github.com/TheWebMonks/meshmonk , v0.0.6), MATLAB 2018a, R (v3.6.1), ggplot2 (v3.1.0), Python(v3.5.0), PLINK v1.9, SHAPEIT2 (v2.17), IMPUTE2 (v2.3.2), SNPLIB (https://github.com/jiarui-li/SNPLIB), GCTA-GREML, FUMA (v1.3.6), GREAT (v4.0.4), Metascape (https://metascape.org), LocusZoom (https://genome.sph.umich.edu/wiki/LocusZoom), and REHH2 (v3.2.0) as mentioned throughout the Methods and supplementary information.

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The meta-analysis GWAS summary statistics are available on the National Omics Data Encyclopedia. NODE: OEP002283 (<https://www.biosino.org/node/project/detail/OEP002283>). The participants making up the NSPT, NHC and TZL datasets were not collected with broad data sharing consent. Given the highly identifiable nature of both facial and genomic information and unresolved issues regarding risk to participants, we opted for a more conservative approach to participant

recruitment. Broad data sharing of the raw data from these collections would thus be in legal and ethical violation of the informed consent obtained from the participants. This restriction is not because of any personal or commercial interests. Additional details can be requested from Li Jin (lijin@fudan.edu.cn) for the NSPT dataset, and Sijia Wang (wangsjia@picb.ac.cn) for the NHC and TZL datasets. Data usage shall be in full compliance with the regulations on management of human genetic resources in China.

Publicly available data used were: 1000 Genomes Project Phase 3: (<https://www.internationalgenome.org/category/phase-3/>), the Roadmap Epigenomics Project: (<http://www.roadmapepigenomics.org>), NCBI dbSNP database: (<http://www.ncbi.nlm.nih.gov/SNP>), HaploReg v4.1: (<https://pubs.broadinstitute.org/mammals/haploreg/haploreg.php>), UCSC genome browser: (<http://genome.ucsc.edu>), GTEx v8: (<https://gtexportal.org/home/>), Ensemble genome browser: (http://asia.ensembl.org/Homo_sapiens/Info/Index), ENCODE: (<https://www.encodeproject.org/>), Human genome dating: (<https://human.genome.dating/>), the transcriptome resource from separated ectoderm and mesenchyme of the developing mouse face (GSE62214). (<https://www.ncbi.nlm.nih.gov/geo/query/acc.cgi?acc=GSE62214>)

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Life sciences study design

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Sample size	The sample size was determined by the number of individuals with genotype data and 3D facial images that were collected with informed consent of three independent studies. The National Survey of Physical Traits (NSPT) cohort (n = 3,322), the Northern Han Chinese (NHC) cohort (n = 4,767), and the Taizhou Longitudinal Study (TZL) cohort (n = 2,881). For the NSPT sample, individuals were recruited at three Chinese cities: Nanning, Guangxi province (n = 1,326); Taizhou, Jiangsu Province (n = 986); Zhengzhou, Henan province (n = 1,010). In the Northern Han Chinese sample, participants were recruited in Tangshan, Hebei province. These two cohorts constituted the discovery dataset. The TZL cohort, where individuals were recruited in Taizhou, Jiangsu province, were used as the replication dataset. A power analysis showed that for the sample size of more than 2000, the statistical power for detecting an additive allele with a standardized effect size of 0.1 at the MAF of 0.2 could achieve more than 80%, and the power would increase up to 95% if the sample size is elevated to 4000. Despite of the fact that the MAFs of some candidate variants were much lower than 0.2, the effect sizes were remarkably large. Accordingly, it seems statistically powerful enough for the sample size in this study.
Data exclusions	Relatedness among the individuals was inferred using KING-robust separately in the discovery and replication datasets. Any relative up to the second-degree was excluded. As a result, 6968 (n = 4,089 in NHC cohort, n = 2,879 in NSPT cohort) and 2706 (in TZL cohort) unrelated individuals in the discovery and replication dataset are used for further analysis. Genotype samples were excluded if there was poor concordance of genetic and reported sex, evidence of chromosomal aberrations, missing genotype call rate > 5%, and heterozygosity values ± 3 standard deviations from the sample mean. For 3D facial images, after GPA and symmetrization, we investigated every mapped image manually and identified outlier images, typically exposed by locally inconsistent triangles on the surface, which are stretched and compressed irregularly in the images. Outlier images can be caused by poor quality or large noise (i.e., isolated pieces, captured position, clothes) of facial images. Four images with poor quality were removed after manually investigated.
Replication	Replication was achieved by proper separation of the data into identification and verification datasets, based on completely separate sampling, imaging, genotyping, and imputation. The TZL cohort, where individuals were recruited in Taizhou, Jiangsu province, were used as the replication dataset. The same phenotype, the top-associated phenotype in the discovery cohort, are used for replication. More information is found in Methods.
Randomization	No randomization took place, group membership of identification and verification was determined by the separately obtained datasets available. Because canonical correlation analysis does not accommodate adjustments for covariates, we removed the effect of relevant covariates (age, age squared, sex, BMI, and four SUGIBS components) using a partial least-squares regression (PLSR, function plsregress from MATLAB 2018a) in both the discovery and replication cohort, and thus performed the canonical correlation analysis under a reduced model with residualized variables.
Blinding	Blinding was not relevant to this study, as no treatment outcomes were assessed and data analysis procedures were standardized across all individuals. Three independent datasets were constructed from the sampling efforts, and the analysis was done by another research team.

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Population characteristics

All participants are Han Chinese. For the NSPT cohort, individuals were recruited at three Chinese cities: Nanning, Guangxi province (n = 1,326); Taizhou, Jiangsu Province (n = 986); Zhengzhou, Henan province (n = 1,010). The percentage of female participants is 63.81%, the average and standard deviation of age is 49.89 and 12.51, the average and standard deviation of BMI is 24.49 and 3.44. For the NHC cohort, the participants were recruited in Tangshan, Hebei province. The percentage of female participants is 51.16%, the average and standard deviation of age is 46.71 and 13.88, the average and standard deviation of BMI is 24.41 and 3.34. The TZL cohort, where individuals were recruited in Taizhou, Jiangsu province, were used as the replication dataset. The percentage of female participants is 64.45%, the average and standard deviation of age is 56.30 and 9.5, the average and standard deviation of BMI is 24.41 and 3.02. For all datasets, there was no imbalance in gender and BMI distributions follow typical distributions seen in an East Asian population. We removed the effect of relevant covariates (age, age squared, sex, BMI, and four SUGIBS components) using a partial least-squares regression (PLSR, function plsregress from MATLAB 2018a) in both the discovery and replication cohort, and thus performed the canonical correlation analysis under a reduced model with residualized variables.

Recruitment

All individual in the three cohorts were imaged using the 2-pod 3dMD Face camera system. Our datasets and the characteristics, as mentioned above, did not contain any specific selection or bias that might influence this work.

Ethics oversight

All participants provided written informed consent, and all study protocols were approved by the institutional review boards of the pertinent research institutions. The Taizhou Longitudinal Study (TZL) was approved by the Ethics Committee of Human Genetic Resources at the Shanghai Institute of Life Sciences, Chinese Academy of Sciences (ER-SIBS-261410). The Northern Han Chinese cohort (NHC) was approved by the Ethics Committee of Human Genetic Resources at the Shanghai Institute of Life Sciences, Chinese Academy of Sciences (ER-SIBS-261410-A1801). The National Survey of Physical Traits (NSPT) is the sub project of The National Science & Technology Basic Research Project which was approved by the Ethics Committee of Human Genetic Resources of School of Life Sciences, Fudan University, Shanghai (14117). We confirm that our study is compliant with the Guidance of the Ministry of Science and Technology (MOST) for the Review and Approval of Human Genetic Resources.

Note that full information on the approval of the study protocol must also be provided in the manuscript.